



Least squares type estimation of the transition density of a particular hidden Markov chain

Claire Lacour

► To cite this version:

Claire Lacour. Least squares type estimation of the transition density of a particular hidden Markov chain. *Electronic Journal of Statistics* , 2008, 2, pp.1-39. hal-00180219

HAL Id: hal-00180219

<https://hal.science/hal-00180219>

Submitted on 18 Oct 2007

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

LEAST SQUARES TYPE ESTIMATION OF THE TRANSITION DENSITY OF A PARTICULAR HIDDEN MARKOV CHAIN

CLAIRE LACOUR

MAP5, Université Paris Descartes, 45 rue des Saints-Pères

75270 Paris Cedex 06, France.

lacour@math-info.univ-paris5.fr

ABSTRACT. In this paper, we study the following model of hidden Markov chain: $Y_i = X_i + \varepsilon_i$, $i = 1, \dots, n+1$ with (X_i) a real-valued stationary Markov chain and $(\varepsilon_i)_{1 \leq i \leq n+1}$ a noise having a known distribution and independent of the sequence (X_i) . We present an adaptive estimator of the transition density obtained by minimization of an original contrast taking advantage of the regressive aspect of the problem. It is selected among a collection of projection estimators with a model selection method. The L^2 -risk and its rate of convergence are evaluated for ordinary smooth noise and some simulations illustrate the method. Our estimator allows to avoid the drawbacks of the quotient estimators.

AMSC 2000: 62G05, 62M05, 62H12

KEYWORDS: Hidden Markov Chain; Transition Density; Nonparametric Estimation; Model Selection; Rate of convergence.

1. INTRODUCTION

In this paper we consider the following additive hidden Markov model:

$$(1) \quad Y_i = X_i + \varepsilon_i \quad i = 1, \dots, n+1$$

with $(X_i)_{i \geq 1}$ a real-valued Markov chain, $(\varepsilon_i)_{i \geq 1}$ a sequence of independent and identically distributed variables and

$$(2) \quad (X_i)_{i \geq 1} \text{ and } (\varepsilon_i)_{i \geq 1} \text{ independent.}$$

Only the variables Y_1, \dots, Y_{n+1} are observed. Besides its initial distribution, the chain $(X_i)_{i \geq 1}$ is characterized by its transition, i.e. the distribution of X_{i+1} given X_i . We assume that this transition has a density Π , defined by $\Pi(x, y)dy = P(X_{i+1} \in dy | X_i = x)$, and our aim is to estimate this transition density Π .

This model belongs to the class of hidden Markov models. The Hidden Markov Models constitute a very famous class of discrete-time stochastic processes, with many applications in various areas such as biology, speech recognition or finance. For a general reference on these models, we refer to Cappé et al. (2005). Here, we study a simple model of HMM where the noise is additive (which allows to deal also with multiplicative noise by use of logarithm). In standard HMM, it is assumed that the joint density of (X_i, Y_i) has a parametric form and the aim is then to infer the parameter from the observations Y_1, \dots, Y_n , generally by maximizing the likelihood. For this type of study, we can see among others Baum and Petrie (1966), Leroux (1992), Bakry et al. (1997), Bickel et al. (1998), Jensen and Petersen (1999), Douc et al. (2004).

This model is also similar to the so-called convolution model (for which the aim is to estimate the density of $(X_i)_{i \geq 1}$). As proceeded for this model, we use extensively the Fourier transform. The restrictions on the error distribution and the rate of convergence obtained for our estimator are also of the same kind. Related works include Stefanski (1990), Fan (1993), Masry (1993) (for the multivariate case), Pensky and Vidakovic (1999), Comte et al. (2006).

The estimation of the transition density of a hidden Markov chain is studied by Cl  men  on (2003). His estimator is based on thresholding of a wavelet-vaguelette decomposition. The drawback of this estimator is that it does not achieve the minimax rate because of a logarithmic loss. Lacour (2007b) describes an estimation procedure by quotient of an estimator of the joint density and an estimator of the stationary density f . The minimax rate is reached by this estimator if we assume that f and Πf have the regularity α . But this smoothness condition on f raises a problem. Indeed Cl  men  on (2000) exhibits an example where the stationary density f is not continuous, whereas the transition density Π is constant. It shows that f can be much less regular than Π . Our aim is then to find an estimator of the transition density which does not have these disadvantages.

To estimate Π , we use an original contrast inspired by the mean square contrast. The first idea is to connect our problem with the regression model. For any function g , we can write

$$g(X_{i+1}) = \left(\int \Pi(\cdot, y)g(y)dy \right) (X_i) + \eta_{i+1}$$

where $\eta_{i+1} = g(X_{i+1}) - \mathbb{E}[g(X_{i+1})|X_i]$. Then, for all function g , we can consider $\int \Pi g$ as a regression function. The mean square contrast to estimate this regression function, if the X_i were known, should be $(1/n) \sum_{i=1}^n [t^2(X_i) - 2t(X_i)g(X_{i+1})]$. If $\int g^2 = 1$, this contrast can be written

$$(1/n) \sum_{i=1}^n \left[\int T^2(X_i, y)dy - 2T(X_i, X_{i+1}) \right]$$

by setting $T(x, y) = t(x)g(y)$ i.e. T such that $\int T(x, y)g(y)dy = t(x)$. It is this contrast which is used in Lacour (2007a) but in our case, only the Y_1, \dots, Y_{n+1} are known. So we introduce in this paper two operators Q and V such that $\mathbb{E}[Q_{T^2}(Y_i)|X_i] = \int T^2(X_i, y)dy$ and $\mathbb{E}[V_T(Y_i, Y_{i+1})|X_i, X_{i+1}] = T(X_i, X_{i+1})$. It leads to the following contrast:

$$(3) \quad \gamma_n(T) = \frac{1}{n} \sum_{i=1}^n [Q_{T^2}(Y_i) - 2V_T(Y_i, Y_{i+1})].$$

A collection of estimators is then defined by minimization of this contrast on wavelet spaces. Indeed wavelets have many useful properties and in particular they can have a compact support and can be regular enough to balance the smoothness of the noise. A general reference on the subject is the book of Meyer (1990).

A method of model selection inspired by Barron et al. (1999) and based on contrast (3) is used to build an adaptive estimator. A data driven choice of model is performed via the minimization of a penalized criterion. The chosen model is the one which minimizes the empirical risk added to a penalty function. In most cases in estimation of mixing processes, a mixing term appears in this penalty. In the same way, some unknown terms derived from the dependence between the X_i appears in the thresholding constant used to define the estimator of Cl  men  on (2003). Here a conditioning argument allows to lead us back to independent variables and thus to avoid such a mixing term in the penalty. Our penalty contains only known quantities or terms that can be estimated and is then computable.

For an ordinary smooth noise with regularity γ , the rate of convergence $n^{-\alpha/(2\alpha+4\gamma+2)}$ is obtained if the transition Π is supposed to belong to a Besov space with regularity α . Our estimator is then better than those of Cl  men  on (2003) which achieves only the rate $(\ln(n)/n)^{\alpha/(2\alpha+4\gamma+2)}$. Moreover this rate is obtained without supposing known the regularity α of f , our estimator is then adaptive.

This paper is organized as follows. In Section 2 we present the model and the assumptions. Section 3 is devoted to the definitions of the contrast and of the estimator. The main result and a sketch of proof are to be found in Section 4. Numerical illustration through simulated examples is reported in Section 5. The detailed proofs are gathered in Section 6.

2. STUDY FRAMEWORK

2.1. Notations. For the sake of clarity, we use lowercase letters for the dimension 1 and capital letters for the dimension 2. For a function $t : \mathbb{R} \mapsto \mathbb{R}$, we denote by $\|t\|$ the L^2 norm that is $\|t\|^2 = \int_{\mathbb{R}} t^2(x)dx$. The Fourier transform t^* of t is defined by

$$t^*(u) = \int e^{-ixu} t(x) dx.$$

Notice that the function t is the inverse Fourier transform of t^* and can be written $t(x) = 1/(2\pi) \int e^{ixu} t^*(u) du$. The convolution product is defined by $(t * s)(x) = \int t(x-y)s(y)dy$.

In the same way, for a function $T : \mathbb{R}^2 \mapsto \mathbb{R}$, $\|T\|^2 = \iint_{\mathbb{R}^2} T^2(x, y) dx dy$ and

$$T^*(u, v) = \iint e^{-ixu - iyv} T(x, y) dx dy.$$

We denote by $t \otimes s$ the function: $(x, y) \mapsto (t \otimes s)(x, y) = t(x)s(y)$.

We will estimate Π on a compact set $A = A_1 \times A_2$ only and we denote by $\|\cdot\|_A$ the norm in $L^2(A)$ i.e.

$$\|T\|_A^2 = \iint_A T^2(x, y) dx dy.$$

2.2. Assumptions on the noise. The Markov chain $(X_i)_{i \geq 1}$ is observed through a noise sequence $(\varepsilon_i)_{i \geq 1}$ of independent and identical distributed random variables. The density of ε_i is denoted by q and is assumed to be known. We assume that the Fourier transform of q never vanishes and that q is ordinary smooth. More precisely the assumption on the error density is the following.

H1 q is uniformly bounded and there exist $\gamma > 0$ and $k_0 > 0$ such that $\forall x \in \mathbb{R} \ |q^*(x)| \geq k_0(x^2 + 1)^{-\gamma/2}$.

This assumption restrains the regularity class of the noise. Among the so-called ordinary smooth noises, we can cite the Laplace distribution, the exponential distribution and all the Gamma or symmetric Gamma distributions. The noise follows a Gamma distribution with scale parameter λ and shape parameter ζ if $q(x) = \lambda^\zeta x^{\zeta-1} e^{-\lambda x} / \Gamma(\zeta)$ for $x > 0$ with Γ the classic Gamma function, and then

$$|q^*(x)| = \left(1 + \frac{x^2}{\lambda^2}\right)^{-\zeta/2}.$$

So q is bounded and verifies H1 with $\gamma = \zeta$. The case $\zeta = 1$ corresponds to an exponential distribution and if $\lambda = 1/2$, $\zeta = p/2$, it is a chi-square $\chi(p)$. A Laplace noise is defined in the following way

$$q(x) = \frac{\lambda}{2} e^{-\lambda|x-\mu|} \text{ and } |q^*(x)| = \frac{\lambda^2}{x^2 + \lambda^2}$$

Then H1 is satisfied with $\gamma = 2$. More generally, we can define the symmetric gamma distribution with density $q(x) = \lambda^\zeta |x|^{\zeta-1} e^{-\lambda|x|} / (2\Gamma(\zeta))$. The characteristic function is then

$$q^*(x) = \left(1 + \frac{x^2}{\lambda^2}\right)^{-\zeta/2} \cos\left(2\zeta \arctan\left(\frac{x}{\lambda + \sqrt{x^2 + \lambda^2}}\right)\right)$$

so that H1 is verified with $\gamma = \zeta + 1$ if ζ is an odd integer and $\gamma = \zeta$ otherwise.

Remark 1. We have to notice that the Gaussian noise does not verify Assumption H1. Indeed, an exponential decrease of the Fourier transform of the error density is more difficult to control and a supersmooth noise makes the denoising more difficult. For that reason, many authors have considered only ordinary smooth noise, one can cite among others Butucea (2004), Koo and Lee (1998) or Youndjé and Wells (2002). The method used in this paper does not allow to deal with

supersmooth noise. Indeed, it requires a basis more regular than the noise and also with compact support (because of Assumption H4 below), which is impossible when the noise is supersmooth.

2.3. Assumptions on the chain. The hypotheses on the hidden Markov chain $(X_i)_{i \geq 1}$ are the following.

H2 The chain is irreducible, positive recurrent and stationary with unknown density f .

H3 There exists a positive real f_0 such that, for all x in A_1 ,

$$f_0 \leq f(x) \leq \|f\|_{\infty, A_1} < \infty$$

H4 The transition density Π is bounded on A by $\|\Pi\|_{\infty, A} < \infty$.

H5 The process (X_k) is geometrically β -mixing ($\beta_q \leq e^{-\theta q}$), or arithmetically β -mixing ($\beta_q \leq q^{-\theta}$) with $\theta > 8$ where

$$\beta_q = \int \|P^q(x, \cdot) - \mu\|_{TV} f(x) dx$$

with $P^q(x, \cdot)$ the distribution of X_{i+q} given $X_i = x$, μ the stationary distribution and $\|\cdot\|_{TV}$ the total variation distance.

We refer to Doukhan (1994) for details on the β -mixing. Assumption H5 implies that the process (Y_k) is β -mixing with β -mixing coefficients smaller than those of (X_k) . Assumption H3 is common (but restrictive) and is crucial to control the empirical processes brought into play. A lot of processes verify Assumptions H2–H5, as autoregressive processes, diffusions or ARCH processes. These examples are detailed in Lacour (2007a).

3. ESTIMATION PROCEDURE

3.1. Projection spaces. Here we describe the projection that we use in this paper to estimate the transition Π . We will consider an increasing sequence of spaces, indexed by m , to construct a collection of estimators. For the sake of simplicity, we assume that $A = [0, 1]^2$.

We use a compactly supported wavelet basis on the interval $[0, 1]$, described in Cohen et al. (1993). The construction furnishes a set of functions (ϕ_k) for $k = 0, \dots, 2^J - 1$ with J a fixed level, and for all $j > J$ a set of functions (ψ_{jk}) , $k = 0, \dots, 2^j - 1$. The collection of these functions forms a complete orthonormal system on $[0, 1]$. Then, for u in $L^2([0, 1])$, we can write

$$u = \sum_{k=0}^{2^J-1} b_k \phi_k + \sum_{j>J} \sum_{k=0}^{2^j-1} a_{jk} \psi_{jk}.$$

Actually

$$\phi_k(x) = \begin{cases} 2^{J/2} \phi^0(2^J x - k) & \text{if } k = 0, \dots, N-1 \\ 2^{J/2} \phi(2^J x - k) & \text{if } k = N, \dots, 2^J - N - 1 \\ 2^{J/2} \phi^1(2^J x - k) & \text{if } k = 2^J - N, \dots, 2^J - 1 \end{cases}$$

where ϕ is a Daubechies father wavelet with support $[-N+1, N]$ and ϕ^0, ϕ^1 are edge wavelets explicitly constructed in Cohen et al. (1993). The functions ϕ_{jk} have support $[(k-N+1)/2^j, (k+N)/2^j] \cap [0, 1]$. For r a positive real, N is chosen large enough so that ϕ has regularity r (it is possible since it is a property of the Daubechies wavelets that the smoothness of ϕ increases linearly with N). We choose J such that $2^J \geq 2N$ so that the two edges do not interact (no overlap between ϕ^0 and ϕ^1). The construction ensures that ϕ^0 and ϕ^1 are also of regularity r . In the same way, for each level j , the ψ_{jk} are dilatation and translation of functions ψ, ψ^0 and ψ^1 with regularity r .

Now we construct a wavelet basis of $L^2([0, 1]^2)$ by the tensorial product method (see Meyer (1990) Chapter 3 Section 3). The father wavelet is $\phi \otimes \phi$ and the mother wavelets are $\phi \otimes \psi, \psi \otimes \phi,$

$\phi \otimes \phi$. A function T in $L^2([0, 1]^2)$ can then be written

$$T(x, y) = \sum_{k=0}^{2^J-1} \sum_{l=0}^{2^J-1} b_{kl} \phi_k(x) \phi_l(y) + \sum_{j>J} \sum_{k=0}^{2^j-1} \sum_{l=0}^{2^j-1} (a_{jkl}^{(1)} \phi_{jk}(x) \psi_{jl}(y) + a_{jkl}^{(2)} \psi_{jk}(x) \phi_{jl}(y) + a_{jkl}^{(3)} \psi_{jk}(x) \psi_{jl}(y)).$$

For the sake of simplicity, we adopt the following notation

$$T(x, y) = \sum_{j \geq J} \sum_{(k,l) \in \Lambda_j} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y).$$

where $\varphi_{jk} = 2^{j/2} \varphi(2^j x - k)$ with $\varphi = \phi, \phi^0, \phi^1, \psi, \psi^0$ or ψ^1 according to the values of j and k . For $j > J$, Λ_j is a set with cardinal $3 \cdot 2^{2j}$ and Λ_J is a set with cardinal 2^{2J} . In the sequel we will use the following property of φ deriving from the regularity of the initial Daubechies wavelet: there exists a positive constant k_3 such that

$$(4) \quad \forall u \in \mathbb{R} \quad |\varphi^*(u)| \leq k_3(u^2 + 1)^{-r/2}$$

Now, for $m \geq J$, we can consider the space

$$\mathbb{S}_m = \{T : \mathbb{R}^2 \rightarrow \mathbb{R}, \quad T(x, y) = \sum_{j=J}^m \sum_{(k,l) \in \Lambda_j} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y)\}$$

Note that the functions in \mathbb{S}_m are all supported in the interval $[0, 1]^2$. The dimension of the space \mathbb{S}_m is $D_m^2 = 2^{2J} + 3 \sum_{j=J+1}^m 2^{2j} \in [2^{2m}, 2^{2m+2}]$. We denote by \mathcal{S} the space \mathbb{S}_{m_0} with the greatest dimension $D_{m_0}^2 = D^2$ smaller than $n^{1/(4\gamma+2)}$. It is the maximal space that we consider. The spaces \mathbb{S}_m have the following properties:

- P1:** $m' \leq m \Rightarrow \mathbb{S}_{m'} \subset \mathbb{S}_m$
- P2:** $\|\sum_{jkl} a_{jkl} \varphi_{jk} \otimes \varphi_{jl}\|^2 = \sum_{jkl} a_{jkl}^2$.

This property derives from the orthonormality of the basis.

Now, for all function $t : \mathbb{R} \mapsto \mathbb{R}$, let v_t be the inverse Fourier transform of $t^*/q^*(-.)$, i.e.

$$v_t(x) = \frac{1}{2\pi} \int e^{ixu} \frac{t^*(u)}{q^*(-u)} du.$$

This operator is introduced because it verifies $\mathbb{E}[v_t(Y_k)|X_k] = t(X_k)$ for all function t . We can write the following lemma :

Lemma 1. *If $r > \gamma + 2$, there exists $\Phi_1 > 0$ such that*

- P3:** $\|\sum_{j=J}^m \sum_k \varphi_{jk}^2\|_\infty \leq \Phi_1 D_m$
- P4:** $\|\sum_k |v_{\varphi_{jk}}|^2\|_\infty \leq \Phi_1 (2^j)^{2\gamma+2}$
- P5:** $\sum_k \|v_{\varphi_{jk}}\|^2 \leq \Phi_1 (2^j)^{2\gamma+1}$
- P6:** $\|\sum_{kk'} |v_{\varphi_{jk} \varphi_{jk'}}|^2\|_\infty \leq \Phi_1 (2^j)^{2\gamma+3}$
- P7:** $\sum_{kk'} \int |v_{\varphi_{jk} \varphi_{jk'}}|^2 \leq \Phi_1 (2^j)^{2\gamma+2}$

This lemma is proved in Section 6.

3.2. Construction of a contrast. We will estimate the transition density of the Markov chain by minimizing a contrast. This section is devoted to the definition of this contrast. We explain here how we can obtain it by considering first the case without noise.

3.2.1. *First step: if the (X_i) 's were observed.* We present here an heuristic to understand why we choose the contrast, assuming first that the (X_i) are known. For all function g , the definition of the transition density implies $\mathbb{E}[g(X_{i+1})|X_i] = \int \Pi(X_i, y) g(y) dy$ so that we can write

$$g(X_{i+1}) = \left(\int \Pi(., y) g(y) dy \right) (X_i) + \eta_i$$

where $\eta_i = g(X_{i+1}) - \mathbb{E}[g(X_{i+1})|X_i]$ is a centered process. We recognize then a regression model. A contrast to estimate $\int \Pi(., y) g(y) dy$ is

$$\gamma_n(u) = \frac{1}{n} \sum_{i=1}^n [u^2(X_i) - 2u(X_i)g(X_{i+1})].$$

It is the classical mean square contrast to estimate a regression function. But we want to estimate $\Pi(., y)$ and not only $\int \Pi(., y) g(y) dy$.

So we observe that if $\int g^2 = 1$ and $T(x, y) = u(x)g(y)$, then $u(.) = \int T(., y)g(y)dy$. So if $u(.) = \int T(., y)g(y)dy$ estimates $\int \Pi(., y)g(y)dy$, we can hope that T estimates Π . Since $\int T^2(., y)dy = u^2(.)$, the contrast becomes

$$\gamma_n(T) = \frac{1}{n} \sum_{i=1}^n [\int T^2(X_i, y) dy - 2T(X_i, X_{i+1})]$$

It is the contrast studied in Lacour (2007a) and it allows a good estimation of $\Pi(., y)$ in the case where the Markov chain is observed. We can observe that

$$\mathbb{E}\gamma_n(T) = \int T^2(x, y)f(x)dx dy - 2 \int T(x, y)f(x)\Pi(x, y)dx dy = \|T - \Pi\|_f^2 - \|\Pi\|_f^2$$

where f is the density of X_i and

$$\|T\|_f = \left(\int T^2(x, y)f(x)dx dy \right)^{1/2}.$$

Then this contrast is an empirical counterpart of the distance $\|T - \Pi\|_f$.

3.2.2. *Second step: the X_i 's are unknown, the observations are the (Y_i) 's.* The aim of this step is to modify the previous contrast, to take into account that the X_i 's are not observed. To do this, we use the same technique as in the convolution problem (see Comte et al. (2006)). Let us denote by F_X the density of (X_i, X_{i+1}) and F_Y the density of (Y_i, Y_{i+1}) . We remark that $F_Y = F_X * (q \otimes q)$ and $F_Y^* = F_X^*(q^* \otimes q^*)$ and then

$$\mathbb{E}[T(X_i, X_{i+1})] = \iint T F_X = \frac{1}{2\pi} \iint T^* \overline{F_X^*} = \frac{1}{2\pi} \iint \frac{T^*}{q^* \otimes q^*} \overline{F_Y^*}$$

by using the Parseval equality. The idea is then to define $V_T^* = T^*/(\overline{q^*} \otimes \overline{q^*})$ so that

$$\mathbb{E}[T(X_i, X_{i+1})] = \frac{1}{2\pi} \iint V_T^* \overline{F_Y^*} = \iint V_T F_Y = \mathbb{E}[V_T(Y_i, Y_{i+1})].$$

Then we will replace the term $T(X_i, X_{i+1})$ in the contrast by $V_T(Y_i, Y_{i+1})$. In the same way, we find an operator Q to replace the term $\int T^2(X_i, y)dy$. More precisely, for all function T , let V_T be the inverse Fourier transform of $T^*/(q^* \otimes q^*)(.)$, i.e.

$$V_T(x, y) = \frac{1}{4\pi^2} \iint e^{ixu+iyv} \frac{T^*(u, v)}{q^*(-u)q^*(-v)} du dv.$$

Let Q_T be the inverse Fourier transform of $T^*(., 0)/(q^*)(.)$, i.e.

$$Q_T(x) = \frac{1}{2\pi} \int e^{ixu} \frac{T^*(u, 0)}{q^*(-u)} du.$$

V and Q have been chosen so that the following lemma holds.

Lemma 2. *For all $k \in \{1, \dots, n+1\}$*

- (1) $\mathbb{E}[V_T(Y_k, Y_{k+1}) | X_1, \dots, X_{n+1}] = T(X_k, X_{k+1})$
- (2) $\mathbb{E}[V_T(Y_k, Y_{k+1})] = \iint T(x, y) \Pi(x, y) f(x) dx dy$
- (3) $\mathbb{E}[Q_T(Y_k) | X_1, \dots, X_{n+1}] = \int T(X_k, y) dy$
- (4) $\mathbb{E}[Q_T(Y_k)] = \iint T(x, y) f(x) dx dy$

Points 1. and 3. are proved in Section 6, the other assertions are immediate consequences. Notice that V and Q are strongly linked with v . In particular $V_{s \otimes t}(x, y) = v_s(x) v_t(y)$ and $Q_{s \otimes t}(x) = v_s(x) \int t(y) dy$.

By using the operators V and Q , we define now the contrast, depending only on the observations Y_1, \dots, Y_{n+1} :

$$\gamma_n(T) = \frac{1}{n} \sum_{k=1}^n [Q_{T^2}(Y_k) - 2V_T(Y_k, Y_{k+1})]$$

With Lemma 2, we compute $\mathbb{E}(\gamma_n(T)) = \iint T^2(x, y) f(x) dx dy - 2 \iint T(x, y) \Pi(x, y) f(x) dx dy = \|T - \Pi\|_f^2 - \|\Pi\|_f^2$. So we want to estimate Π by minimizing γ_n . The definition of the contrast leads to the following “empirical norm”:

$$\Psi_n(T) = \frac{1}{n} \sum_{k=1}^n Q_{T^2}(Y_k).$$

The term empirical norm is used because $\mathbb{E}\Psi_n(T) = \|T\|_f^2$, but Ψ_n is not a norm in the common sense.

3.3. Definition of the estimator. We have to minimize the contrast γ_n to find our estimator.

By writing $T = \sum_{j=J}^m \sum_{(k,l) \in \Lambda_j} a_{jkl} \varphi_{jk} \otimes \varphi_{jl} = \sum_{\lambda} a_{\lambda} \omega_{\lambda}(x, y)$, we obtain

$$\frac{\partial \gamma_n(T)}{\partial a_{\lambda_0}} = \frac{2}{n} \sum_{i=1}^n \left(\sum_{\lambda} a_{\lambda} Q_{\omega_{\lambda} \omega_{\lambda_0}}(Y_i) - V_{\omega_{\lambda_0}}(Y_i, Y_{i+1}) \right).$$

Then, by denoting by A_m the matrix of the coefficients a_{λ} of T ,

$$(5) \quad \forall \lambda_0 \quad \frac{\partial \gamma_n(t)}{\partial a_{\lambda_0}} = 0 \quad \Longleftrightarrow \quad G_m A_m = Z_m$$

where

$$G_m = \left[\frac{1}{n} \sum_{i=1}^n Q_{\omega_{\lambda} \omega_{\mu}}(Y_i) \right]_{\lambda, \mu}, \quad Z_m = \left[\frac{1}{n} \sum_{i=1}^n V_{\omega_{\lambda}}(Y_i, Y_{i+1}) \right]_{\lambda}$$

But the matrix G_m is not necessarily invertible. That is why we introduce the set

$$(6) \quad \Gamma = \left\{ \min \text{Sp}(G_m) \geq \frac{2}{3} f_0 \right\}$$

where Sp denotes the spectrum, i.e. the set of the eigenvalues of the matrix and f_0 is the lower bound of f on A_1 . On Γ , G_m is invertible and γ_n is convex so that the minimization of γ_n is equivalent to Equation (5) and admits the solution $A_m = G_m^{-1} Z_m$. Now we can define

$$\hat{\Pi}_m = \begin{cases} \arg \min_{T \in \mathbb{S}_m} \gamma_n(T) & \text{on } \Gamma, \\ 0 & \text{on } \Gamma^c. \end{cases}$$

Remark 2. The construction of $\hat{\Pi}_m$ described here requires the knowledge of f_0 . Nevertheless, when f_0 is unknown, we can replace it by an estimator \hat{f}_0 defined as the minimum of an estimator of f (for an estimator of the density of a hidden Markov chain, see Lacour (2007b)). The result is then unchanged if f is regular enough and the mixing rate high enough.

We have then an estimator of Π for all \mathbb{S}_m . But we have to choose the best model m to obtain an adaptive estimator, i.e. an estimator which achieves the best rate of convergence whatever the regularity of Π . So we set

$$\hat{m} = \arg \min_{m \in \mathcal{M}_n} \{\gamma_n(\hat{\Pi}_m) + \text{pen}(m)\}$$

where pen is a penalty function to be specified later and

$$\mathcal{M}_n = \{m \geq J, D_m^{4\gamma+2} \leq n\}.$$

Then we can define our definitive estimator:

$$\tilde{\Pi} = \begin{cases} \hat{\Pi}_{\hat{m}} & \text{if } \|\hat{\Pi}_{\hat{m}}\| \leq k_n \quad \text{with } k_n = n^{1/2}, \\ 0 & \text{else.} \end{cases}$$

4. RESULT

4.1. Risk and rate of convergence. For a function G and a subspace \mathbb{S} , we define

$$d(G, \mathbb{S}) = \inf_{T \in \mathbb{S}} \|G - T\|_A.$$

We recall that A is the estimation area. For each estimator $\hat{\Pi}_m$, we have the following decomposition of the risk.

Proposition 1. We consider a Markov chain and a noise satisfying Assumptions H1–H5 with $\gamma \geq 3/4$. For m fixed in \mathcal{M}_n , we consider $\hat{\Pi}_m$ the estimator of the transition density Π previously described. Then there exists $C > 0$ such that

$$\mathbb{E}\|\hat{\Pi}_m - \Pi\|_A^2 \leq C \left\{ d_A^2(\Pi, S_m) + \frac{D_m^{4\gamma+2}}{n} \right\}.$$

We do not prove this proposition because this result is included in Theorem 1 below, which is proved in Section 6.

Now if Π belongs to a Besov space with regularity α , it is a common approximation property of the wavelet spaces that $d^2(\Pi, S_m) \leq CD_m^{-2\alpha}$. So, choosing m_1 such that $D_{m_1} = n^{1/(2\alpha+4\gamma+2)}$, we can obtain the minimum risk

$$\mathbb{E}\|\hat{\Pi}_{m_1} - \Pi\|_A^2 \leq Cn^{-\frac{2\alpha}{2\alpha+4\gamma+2}}.$$

But this choice of m_1 is impossible if α is unknown (it is *a priori* the case since Π is unknown). That is why we have built an adaptive estimator via model selection. For our estimator $\tilde{\Pi}$, we can state the following theorem.

Theorem 1. We consider a Markov chain and a noise satisfying Assumptions H1–H5 with $\gamma > 3/4$. We consider $\tilde{\Pi}$ the estimator of the transition density Π previously described with $r > 2\gamma + 3/2$ and

$$\text{pen}(m) = K \frac{D_m^{4\gamma+2}}{n} \text{ for some } K > K_0$$

where K_0 is a constant depending on Φ_1 , $\|q\|_\infty$ and f_0 . Then there exists $C' > 0$ such that

$$\mathbb{E}\|\tilde{\Pi} - \Pi\|_A^2 \leq C \inf_{m \in \mathcal{M}_n} \{d_A^2(\Pi, S_m) + \text{pen}(m)\} + \frac{C'}{n}$$

with $C = \max(2 + 72f_0^{-1}\|f\|_{\infty, A_1}(1 + 2\|\Pi\|_A^2), 12f_0^{-1}(1 + 2\|\Pi\|_A^2))$.

Note that this result is non-asymptotic, it is an advantage of the least square method with respect to a quotient method.

All constants on which the penalty depends do not have the same status. The constants Φ_1 and $\|q\|_\infty$ are known, since the wavelet basis and the noise distribution are known. The constant f_0 is unknown but it can be estimated (see Remark 2). Then, even if it means replacing f_0 by an estimator \hat{f}_0 , the penalty is computable. In particular the dependence coefficients of the sequence do not appear at all in the penalty.

The condition $\gamma > 3/4$ is due to an additional term of order $D_m^{2\gamma+7/2}/n$ (coming from the term $(1/n)\sum_{i=1}^n Q_{T^2}(Y_i)$ in the contrast) inside the penalty. If $\gamma > 3/4$, then $2\gamma + 7/2 < 4\gamma + 2$ and $D_m^{4\gamma+2}/n$ is the dominant term. If $\gamma = 3/4$, the result is still true but the constant in the penalty also depends on $\|\Pi\|_A$. In the other cases the estimation is possible but the term $D_m^{2\gamma+4}/n$ is not negligible any more and the order of the variance (and consequently the rate of convergence) must be changed. This constraint $\gamma > 3/4$ is not restrictive since γ must be larger than $1/2$ in order that q be square integrable. Moreover in the case of a Gamma noise, q is not bounded if $\gamma < 1$.

We can now evaluate the rate of convergence of our estimator.

Corollary 1. *We suppose that the restriction of Π to A belongs to the Besov space $B_{2,\infty}^\alpha(A)$ with $\alpha < r$. Then, under the assumptions of Theorem 1,*

$$\mathbb{E}\|\tilde{\Pi} - \Pi\|_A^2 = O(n^{-\frac{2\alpha}{2\alpha+4\gamma+2}}).$$

There is to our knowledge no lower bound for our precise estimation problem and such a study is beyond the scope of the present work. But Cl  men  on (2003) proves the optimality of the rate $n^{-\frac{2\alpha}{2\alpha+4\gamma+2}}$ in the case where f belongs to $B_{2,\infty}^\alpha(\mathbb{R})$ and $f\Pi$ belong to $B_{2,\infty}^\alpha(\mathbb{R}^2)$.

Nevertheless we remark that we obtain then the same rate of convergence with $\tilde{\Pi}$ as those obtained with $\hat{\Pi}_{m_1}$ where $D_{m_1} = n^{1/(2+4\gamma+2\alpha)}$, but without requiring the knowledge of α . That is why we can assert that this estimator is adaptive. Moreover our estimator is better than the one of Cl  men  on (2003), which achieves only the rate $(\ln(n)/n)^{\frac{2\alpha}{2\alpha+4\gamma+2}}$. It is also an improvement of the result of Lacour (2007b) because this rate is obtained without requiring that f has a regularity α .

4.2. Sketch of proof of Theorem 1. We give in this section a sketch of proof of Theorem 1.

Let $m \in \mathcal{M}_n$. We denote by Π_m the orthogonal projection of Π on \mathbb{S}_m . We have the following bias-variance decomposition

$$\mathbb{E}\|\tilde{\Pi} - \Pi\|_A^2 = \mathbb{E}\|\tilde{\Pi} - \Pi_m\|_A^2 + \|\Pi_m - \Pi\|_A^2$$

The term $\|\tilde{\Pi} - \Pi_m\|_A^2$ can be written in the following way

$$\begin{aligned} \|\tilde{\Pi} - \Pi_m\|_A^2 &= \|\tilde{\Pi} - \Pi_m\|_A^2 \mathbf{1}_{\{\|\hat{\Pi}_{\hat{m}}\| \leq k_n\}} + \|\tilde{\Pi} - \Pi_m\|_A^2 \mathbf{1}_{\{\|\hat{\Pi}_{\hat{m}}\| > k_n\}} \\ &\leq \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_A^2 + \|\Pi_m\|_A^2 \mathbf{1}_{\{\|\hat{\Pi}_{\hat{m}}\| > k_n\}} \end{aligned}$$

since $\tilde{\Pi} = 0$ on the set $\{\|\hat{\Pi}_{\hat{m}}\| > k_n\}$ and $\tilde{\Pi} = \hat{\Pi}_{\hat{m}}$ on the complement. The term $\|\Pi_m\|_A^2 \mathbf{1}_{\{\|\hat{\Pi}_{\hat{m}}\| > k_n\}}$ is easy to deal with, the main term is $\|\hat{\Pi}_{\hat{m}} - \Pi_m\|_A^2$. But, on Γ , the definitions of $\hat{\Pi}_m$ and \hat{m} lead to the inequality

$$(7) \quad \gamma_n(\hat{\Pi}_{\hat{m}}) + \text{pen}(\hat{m}) \leq \gamma_n(\Pi_m) + \text{pen}(m).$$

Letting $Z_{n,m}(T) = \frac{1}{n} \sum_{k=1}^n [V_T(Y_k, Y_{k+1}) - Q_{T\Pi_m}(Y_k)]$, a fast computation gives

$$\gamma_n(\hat{\Pi}_{\hat{m}}) - \gamma_n(\Pi_m) = \Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m) - 2Z_{n,m}(\hat{\Pi}_{\hat{m}} - \Pi_m)$$

so that (7) becomes

$$\begin{aligned}\Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m) &\leq 2Z_{n,m}(\hat{\Pi}_{\hat{m}} - \Pi_m) + \text{pen}(m) - \text{pen}(\hat{m}) \\ &\leq 2\|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f \sup_{T \in B_f(m, \hat{m})} Z_{n,m}(T) + \text{pen}(m) - \text{pen}(\hat{m})\end{aligned}$$

where $B_f(m, \hat{m}) = \{T \in S_m + S_{\hat{m}}, \|T\|_f = 1\}$. The main steps of the proof are then

- (1) to control the term $\sup_{T \in B_f(m, \hat{m})} Z_{n,m}(T)$,
- (2) to link the empirical “norm” Ψ_n with the L^2 norm $\|\cdot\|_A$.

• To deal with the supremum of the empirical process $Z_{n,m}(T)$, we will use an inequality of Talagrand stated in Lemma 5 (Section 6.8). This inequality is very powerful but can be applied only to sum of independent random variables. That is why we split $Z_{n,m}(T)$ into three processes plus a bias term.

$$Z_{n,m}(T) = Z_n^{(1)}(T) - Z_n^{(2)}(T) + Z_n^{(3)}(T) + \iint T(x, y)(\Pi - \Pi_m)(x, y)f(x)dx dy$$

with

$$\begin{cases} Z_n^{(1)}(T) = \frac{1}{n} \sum_{k=1}^n V_T(Y_k, Y_{k+1}) - \mathbb{E}[V_T(Y_k, Y_{k+1}) | X_1, \dots, X_{n+1}] \\ Z_n^{(2)}(T) = \frac{1}{n} \sum_{k=1}^n Q_{T\Pi_m}(Y_k) - \mathbb{E}[Q_{T\Pi_m}(Y_k)] \\ Z_n^{(3)}(T) = \frac{1}{n} \sum_{k=1}^n T(X_k, X_{k+1}) - \mathbb{E}[T(X_k, X_{k+1})] \end{cases}$$

For the first process $Z_n^{(1)}$, we get back to independent variables by remarking that, conditionally to X_1, \dots, X_{n+1} , the couples (Y_{2i-1}, Y_{2i}) are independent (see Proposition 3).

For the other processes, we use the mixing assumption H5 to build auxiliary variables X_i^* which are approximation of the X_i ’s and which constitute independent clusters of variables (see Proposition 4).

- To pass from Ψ_n to the L^2 norm, we introduce the following set

$$\Delta = \{\forall T \in \mathcal{S} \quad \|T\|_f^2 \leq \frac{3}{2}\Psi_n(T)\}$$

We can easily prove (see Section 6.3) that $\Delta \subset \Gamma$. Then,

$$\|\hat{\Pi}_{\hat{m}} - \Pi_m\|_A \mathbf{1}_\Delta \leq \frac{3}{2}f_0^{-1}\Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m)\mathbf{1}_\Gamma$$

It remains to prove that $P(\Delta^c) = P(\exists T \in \mathcal{S}, \Psi_n(T) < (2/3)\mathbb{E}[\Psi_n(T)])$ is small enough. It is done in Proposition 2.

5. SIMULATIONS

To illustrate the method, we compute our estimator $\tilde{\Pi}$ for different Markov processes with known transition density. The estimation procedure contains several Fourier transforms, which can seem heavy, but the computation of $v_{\varphi_{jk}}$ for all the basis functions can be done beforehand. Next, to compute $\tilde{\Pi}$ from data Y_1, \dots, Y_{n+1} , we use the following steps (see Section 3.3):

- For each m , compute matrices G_m and Z_m ,
- Deduce the matrix A_m ,
- Select the \hat{m} which minimize $\gamma_n(\hat{\Pi}_m) + \text{pen}(m) = -{}^t A_m Z_m + \text{pen}(m)$,
- Compute $\tilde{\Pi}$ using matrix $A_{\hat{m}}$.

We consider several kinds of Markov chains :

- An autoregressive process denoted by AR and defined by:

$$X_{n+1} = aX_n + b + \varepsilon_{n+1}$$

where the ε_{n+1} are independent and identical distributed random variables, with centered Gaussian distribution with variance σ^2 . For this process, the transition density can be written $1/(\sigma\sqrt{2\pi}) \exp(-(y - ax - b)^2/2\sigma^2)$. We consider the following parameter values :

- (i) $a = 2/3$, $b = 0$, $\sigma^2 = 5/9$, estimated on $[-2, 2]^2$.
- (ii) $a = 0.5$, $b = 3$, $\sigma^2 = 1$, and then the process is estimated on $[4, 8]^2$.
- A radial Ornstein-Uhlenbeck process (in its discrete version). For $j = 1, \dots, \delta$, we define the processes: $\xi_{n+1}^j = a\xi_n^j + \beta\varepsilon_n^j$ where the ε_n^j are i.i.d. standard Gaussian. The chain is then defined by $X_n = \sqrt{\sum_{i=1}^{\delta} (\xi_n^i)^2}$. The transition density is given in Chaleyat-Maurel and Genon-Catalot (2006) where this process is studied in detail:

$$\pi(x, y) = \mathbb{1}_{y>0} \exp\left(-\frac{y^2 + a^2x^2}{2\beta^2}\right) I_{\delta/2-1}\left(\frac{axy}{\beta^2}\right) \frac{ax}{\beta^2} \left(\frac{y}{ax}\right)^{\delta/2}$$

and $I_{\delta/2-1}$ is the Bessel function with index $\delta/2 - 1$. This process (with here $a = 0.5$, $\beta = 3$, $\delta = 3$) is denoted by $\sqrt{\text{CIR}}$ since its square is actually a Cox-Ingersoll-Ross process. The estimation domain for this process is $[2, 10]^2$.

- A Cox-Ingersoll-Ross process, which is exactly the square of the previous process. It follows a Gamma density for invariant distribution with scale parameter $l = 1/2\rho^2$ and shape parameter $a = \delta/2$. The transition density is

$$\pi(x, y) = \frac{1}{2\beta^2} \exp\left(-\frac{y + a^2x}{2\beta^2}\right) I_{\delta/2-1}\left(\frac{a\sqrt{xy}}{\beta^2}\right) \left(\frac{y}{a^2x}\right)^{\delta/4-1/2}$$

The used parameters are the following:

- (iii) $a = 3/4$, $b = \sqrt{7/48}$ and $\delta = 4$, estimated on $[0.1, 3]^2$.
- (iv) $a = 1/3$, $b = 3/4$ and $\delta = 2$. This chain is estimated on $[0, 2]^2$.
- An ARCH process defined by $X_{n+1} = \sin(X_n) + (\cos(X_n) + 3)\varepsilon_{n+1}$ where the ε_{n+1} are i.i.d. standard Gaussian. The transition density of this chain is

$$\pi(x, y) = \varphi\left(\frac{y - \sin(x)}{\cos(x) + 3}\right) \frac{1}{\cos(x) + 3}$$

and we estimate this process on $[-5, 5]^2$.

For this last chain, the stationary density is not explicit. So we simulate $n + 500$ variables and we estimate only from the last n to ensure the stationarity of the process. For the other chains, it is sufficient to simulate an initial variable X_0 with density f .

We consider two different noises:

Laplace noise: In this case, the density of ε_i is given by

$$q(x) = \frac{\lambda}{2} e^{-\lambda|x|}; \quad q^*(x) = \frac{\lambda^2}{\lambda^2 + x^2}; \quad \lambda = 5.$$

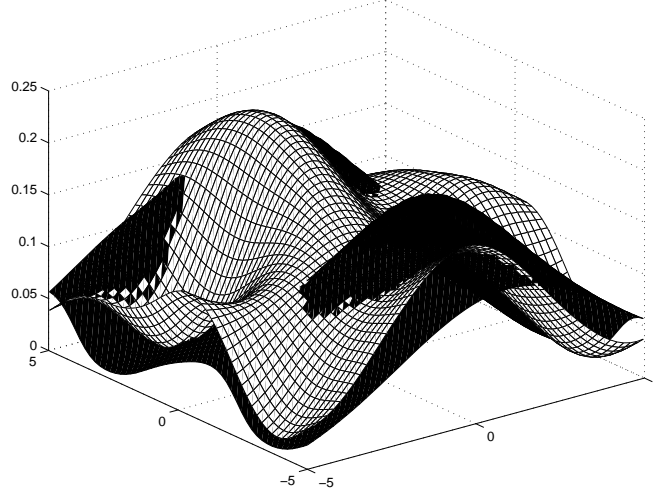
The smoothness parameter is $\gamma = 2$ and we choose

$$\text{pen}(m) = \frac{1}{n} \left(\frac{\lambda}{2}\right)^2 \left(\frac{D_m}{4}\right)^{10}.$$

Gaussian noise: In this case, the density of ε_i is given by

$$q(x) = \frac{1}{\lambda\sqrt{2\pi}} e^{-\frac{x^2}{2\lambda^2}}; \quad q^*(x) = e^{-\frac{\lambda^2 x^2}{2}}; \quad \lambda = 0.3.$$

n	50	100	250	500	1000	noise
AR(i)	0.579	0.407	0.270	0.230	0.209	Lapl
	0.599	0.480	0.313	0.272	0.245	Gauss
AR(ii)	0.389	0.294	0.195	0.155	0.139	Lapl
	0.339	0.304	0.280	0.273	0.271	Gauss
$\sqrt{\text{CIR}}$	0.171	0.138	0.123	0.118	0.111	Lapl
	0.199	0.169	0.150	0.142	0.139	Gauss
CIR(iii)	0.420	0.345	0.237	0.195	0.175	Lapl
	0.337	0.302	0.276	0.245	0.209	Gauss
CIR(iv)	0.525	0.403	0.337	0.304	0.292	Lapl
	0.369	0.345	0.344	0.327	0.321	Gauss
ARCH	0.312	0.287	0.261	0.185	0.150	Lapl
	0.337	0.319	0.296	0.290	0.183	Gauss

TABLE 1. MISE $\mathbb{E}\|\Pi - \tilde{\Pi}\|^2$ averaged over $N = 200$ samples.FIGURE 1. True Π (black) and estimator $\tilde{\Pi}$ (white) for process ARCH observed through a Laplace noise, $n = 500$

This noise does not verify Assumption H1 but it is interesting to see if this assumption is also necessary for practical purposes. Given the exponential regularity of this noise, we consider the following penalty

$$\text{pen}(m) = \frac{5}{n} \exp(\lambda^2 D_m^2).$$

Table 1 presents the L^2 risk of our estimator of the transition density for the 6 Markov chains and the 2 noises. These results can be compared with those of Lacour (2007a) (Table 2) who studies the processes AR(i), $\sqrt{\text{CIR}}$ and ARCH but directly observed, i.e. without noise. The risk values are then higher in our case, but with the same order, which is satisfactory. It is noticeable that the estimation works pretty much the same with the Gaussian noise, but with a slower decrease of the risk. It is a classical phenomenon in deconvolution problems, since the Gaussian noise is much more regular than the Laplace noise.

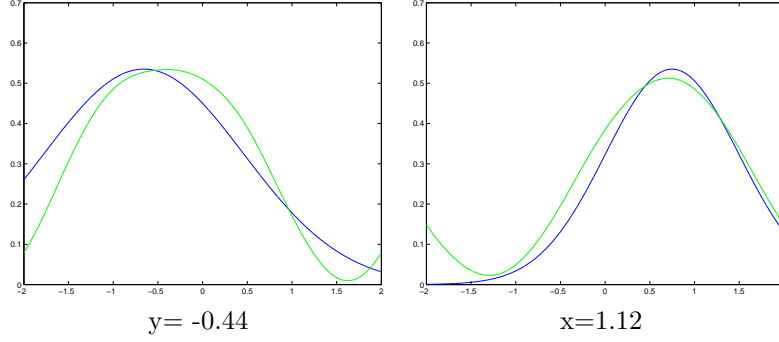
FIGURE 2. Sections for the process AR(i) observed through a Laplace noise, $n = 500$

Figure 5 allows to visualize the result for the process ARCH observed through a Laplace noise: the surfaces $z = \Pi(x, y)$ and $z = \tilde{\Pi}(x, y)$ are presented. We also give figures of sections of this kind of surfaces. We can see on Figure 5 the curves $z = \Pi(x, -0.44)$ versus $z = \tilde{\Pi}(x, -0.44)$ and the curves $z = \Pi(1.12, y)$ versus $z = \tilde{\Pi}(1.12, y)$ for the process AR(i). Generally, for a multidimensional estimation, the mixed control of the directions does not allow to do as well as a classical one-dimensionnal function estimation. Nevertheless the curves are here very close.

6. DETAILED PROOFS

6.1. Proof of Lemma 1. • Using that

$$(8) \quad \left| \sum_k \varphi_{jk}^2(x) \right| \leq C(\varphi) 2^{j/2} \|\varphi_{jk}\|_\infty \leq C'(\varphi) 2^j,$$

P1 holds if $\Phi_1 \geq 2C'(\varphi)$.

- The computation of $v_{\varphi_{jk}}$ gives

$$|v_{\varphi_{jk}}(x)| \leq \frac{2^{j/2}}{2\pi} \int \frac{|\varphi^*(v)|}{|q^*(-v2^j)|} dv$$

Next, it follows from Assumption H1 that $|v_{\varphi_{jk}}(x)| \leq C_{1,\gamma} (2^j)^{\gamma+1/2} / 2\pi k_0$ using Lemma 4 since $r > \gamma + 1$. Then, for all x , $\sum_k |v_{\varphi_{jk}}(x)|^2 \leq 3 \cdot 2^j C_{1,\gamma}^2 \frac{k_0^{-2}}{4\pi^2} (2^j)^{2\gamma+1}$ that establishes P4 with $\Phi_1 \geq 3C_{1,\gamma}^2 k_0^{-2} / (4\pi^2)$.

- To prove P5, we apply the Parseval equality. That yields

$$\int |v_{\varphi_{jk}}|^2 = \frac{1}{2\pi} \int \frac{|\varphi^*(v)|^2}{|q^*(-v2^j)|^2} dv.$$

Using H1 and given that $2r > 2\gamma + 1$, we obtain $\int |v_{\varphi_{jk}}|^2 \leq C_{2,2\gamma} (2^j)^{2\gamma} / 2\pi k_0^2$. And finally P5 holds with $\Phi_1 \geq 3C_{2,2\gamma} k_0^{-2} / (2\pi)$.

- We begin with computing $|v_{\varphi_{jk}\varphi_{jk'}}(x)|$ by using that $(\varphi_{jk}\varphi_{jk'})^*$ is equal to the convolution product $\varphi_{jk}^* * \varphi_{jk'}^*$.

$$\begin{aligned} |v_{\varphi_{jk}\varphi_{jk'}}(x)| &\leq \frac{|\varphi^*(v/2^j)| |\varphi^*((u-v)/2^j)|}{|q^*(-u)|} dudv \\ &\leq \frac{k_0^{-1}}{2\pi} 2^j (2^j)^\gamma \iint |\varphi^*(y)\varphi^*(x-y)| (x^2+1)^{\gamma/2} dx dy. \end{aligned}$$

Then Lemma 4 (Section 6.8) shows that

$$|v_{\varphi_{jk}\varphi_{jk'}}(x)| \leq \frac{k_0^{-1}}{2\pi}(2^j)^{\gamma+1}C_r \left[\int_{|x|>1} |x|^{1-r}(x^2+1)^{\gamma/2} dx + \int_{|x|\leq 1} (x^2+1)^{\gamma/2} dx \right].$$

Hence, since $r > \gamma + 2$, there exists $C > 0$ such that $|v_{\varphi_{jk}\varphi_{jk'}}(x)| \leq C(2^j)^{\gamma+1}$. The fact that φ_{jk} and $\varphi_{jk'}$ have disjoint supports if $k + N \leq k' - N + 1$ or $k' + N \leq k - N + 1$ allows to prove P6 with $\Phi_1 \geq 3(4N - 3)C^2$.

- Applying Parseval's equality,

$$\int |v_{\varphi_{jk}\varphi_{jk'}}|^2 = \frac{2^j}{2\pi} \int \frac{|(\varphi_{jk}\varphi_{jk'})^*(2^j v)|^2}{|q^*(-2^j v)|^2} dv.$$

But, using Lemma 4,

$$(9) \quad |(\varphi_{jk}\varphi_{jk'})^*(2^j v)| \leq \int |\varphi^*(y)| |\varphi^*(v-y)| dy \leq C_r [|v|^{1-r} \mathbf{1}_{|v|>1} + \mathbf{1}_{|v|\leq 1}]$$

Then, it follows that

$$\int |v_{\varphi_{jk}\varphi_{jk'}}|^2 \leq \frac{k_0^{-2} 2^j}{2\pi} C_r^2 \int (|v|^{2(1-r)} \mathbf{1}_{|v|>1} + \mathbf{1}_{|v|\leq 1}) ((2^j v)^2 + 1)^\gamma dv \leq C(2^j)^{2\gamma+1}.$$

It is then sufficient to sum this quantity for all k, k' by taking account of the superposition of the supports to prove P7 as soon as $\Phi_1 \geq 3C(4N - 3)$.

6.2. Proof of Lemma 2. First we write that

$$V_T(Y_k, Y_{k+1}) = \frac{1}{4\pi^2} \int e^{iY_k u + iY_{k+1} v} \frac{T^*(u, v)}{q^*(-u)q^*(-v)} dudv$$

so that, by denoting $\mathbb{X} = (X_1, \dots, X_{n+1})$,

$$\mathbb{E}[V_T(Y_k, Y_{k+1})|\mathbb{X}] = \frac{1}{4\pi^2} \int \mathbb{E}[e^{iY_k u + iY_{k+1} v}|\mathbb{X}] \frac{T^*(u, v)}{q^*(-u)q^*(-v)} dudv.$$

By using the independence between (X_i) and (ε_i) , we compute

$$\begin{aligned} \mathbb{E}[e^{iY_k u + iY_{k+1} v}|\mathbb{X}] &= \mathbb{E}[e^{iX_k u + iX_{k+1} v} e^{i\varepsilon_k u} e^{i\varepsilon_{k+1} v}|\mathbb{X}] \\ &= e^{iX_k u + iX_{k+1} v} \mathbb{E}[e^{i\varepsilon_k u}] \mathbb{E}[e^{i\varepsilon_{k+1} v}] = e^{iX_k u + iX_{k+1} v} q^*(-u)q^*(-v). \end{aligned}$$

Then

$$\mathbb{E}[V_T(Y_k, Y_{k+1})|\mathbb{X}] = \frac{1}{4\pi^2} \int e^{iX_k u + iX_{k+1} v} T^*(u, v) dudv = T(X_k, X_{k+1}).$$

We proceed in a similar way for Q . Since $Q_T(Y_k) = (1/2\pi) \int e^{iY_k u} T^*(u, 0) (q^*(-u))^{-1} du$, then

$$\mathbb{E}[Q_T(Y_k)|\mathbb{X}] = \frac{1}{2\pi} \int \mathbb{E}[e^{iY_k u}|\mathbb{X}] \frac{T^*(u, 0)}{q^*(-u)} du.$$

By using the independence between (X_i) and (ε_i) , we compute

$$\mathbb{E}[e^{iY_k u}|\mathbb{X}] = \mathbb{E}[e^{iX_k u} e^{i\varepsilon_k u}|\mathbb{X}] = e^{iX_k u} \mathbb{E}[e^{i\varepsilon_k u}] = e^{iX_k u} q^*(-u).$$

Thus

$$\mathbb{E}[Q_T(Y_k)|\mathbb{X}] = \frac{1}{2\pi} \int e^{iX_k u} q^*(-u) \frac{T^*(u, 0)}{q^*(-u)} du = \frac{1}{2\pi} \int e^{iX_k u} T^*(u, 0) du.$$

By denoting by T_y the function $x \mapsto T_y(x) = T(x, y)$, we obtain

$$T^*(u, 0) = \iint e^{-ixu} T_y(x) dx dy = \int T_y^*(u) dy$$

and then

$$(10) \quad \begin{aligned} \frac{1}{2\pi} \int e^{iX_k u} T^*(u, 0) du &= \frac{1}{2\pi} \iint e^{iX_k u} T_y^*(u) dy du \\ &= \int T_y(X_k) dy = \int T(X_k, y) dy. \end{aligned}$$

6.3. Proof of Theorem 1. We start with introducing some auxiliary variables whose existence is ensured by Assumption H5 of mixing. In the case of arithmetical mixing, since $\theta > 8$, there exists a real c such that $0 < c < 3/8$ and $c\theta > 3$. We set in this case $q_n = \lfloor n^c \rfloor$. In the case of geometrical mixing, we set $q_n = \lfloor c \ln(n) \rfloor$ where c is a real larger than $3/\theta$.

For the sake of simplicity, we suppose that $n + 1 = 2p_n q_n$, with p_n an integer. Let for $l = 0, \dots, p_n - 1$, $A_l = (X_{2lq_n+1}, \dots, X_{(2l+1)q_n})$, $B_l = (X_{(2l+1)q_n+1}, \dots, X_{(2l+2)q_n})$. As in Viennet (1997), by using Berbee's coupling Lemma, we can build a sequence (A_l^*) such that

$$(11) \quad \begin{cases} A_l \text{ and } A_l^* \text{ have the same distribution,} \\ A_l^* \text{ and } A_{l'}^* \text{ are independent if } l \neq l', \\ P(A_l \neq A_l^*) \leq \beta_{q_n}. \end{cases}$$

In the same way, we build (B_l^*) and we define for any $l \in \{0, \dots, p_n - 1\}$, $A_l^* = (X_{2lq_n+1}^*, \dots, X_{(2l+1)q_n}^*)$, $B_l^* = (X_{(2l+1)q_n+1}^*, \dots, X_{(2l+2)q_n}^*)$ so that the sequence (X_1^*, \dots, X_n^*) is well defined. We can now define

$$\Omega_X^* = \{\forall i, 1 \leq i \leq n+1 \quad X_i = X_i^*\}.$$

Let us recall that \mathcal{S} is the space \mathbb{S}_m with maximal dimension $\mathcal{D}^2 \leq n^{\frac{1}{4\gamma+2}}$. We now adopt the notations

$$\Delta = \{\forall T \in \mathcal{S} \quad \|T\|_f^2 \leq \frac{3}{2} \Psi_n(T)\}; \quad \Omega = \Delta \cap \Omega_X^*.$$

Let us fix $m \in \mathcal{M}_n$. We denote by Π_m the orthogonal projection of Π on \mathbb{S}_m . Then we have the decomposition

$$\begin{aligned} \mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 &\leq 2\mathbb{E} \left(\|\tilde{\Pi} - \Pi_m\|^2 \mathbf{1}_{\|\hat{\Pi}_{\hat{m}}\| \leq k_n} \right) + 2\mathbb{E} \left(\|\tilde{\Pi} - \Pi_m\|^2 \mathbf{1}_{\|\hat{\Pi}_{\hat{m}}\| > k_n} \right) \\ &\quad + 2\mathbb{E} \left(\|\tilde{\Pi} - \Pi_m\|^2 \mathbf{1}_{\Omega^c} \right) + 2\|\Pi_m - \Pi\|_A^2 \\ &\leq 2\mathbb{E} \left(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_{\Omega} \right) + 2\|\Pi_m\|^2 \mathbb{E} \left(\mathbf{1}_{\Omega} \mathbf{1}_{\|\hat{\Pi}_{\hat{m}}\| > k_n} \right) \\ &\quad + 2\mathbb{E} \left([2\|\tilde{\Pi}\| + 2\|\Pi_m\|^2] \mathbf{1}_{\Omega^c} \right) + 2\|\Pi_m - \Pi\|_A^2. \end{aligned}$$

Now, using the Markov inequality and the definition of $\tilde{\Pi}$,

$$\begin{aligned} \mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 &\leq 2\mathbb{E} \left(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_{\Omega} \right) + 2\|\Pi\|_A^2 \frac{\mathbb{E}(\|\hat{\Pi}_{\hat{m}}\|^2 \mathbf{1}_{\Omega})}{k_n^2} \\ &\quad + 4(k_n^2 + \|\Pi\|_A^2) \mathbb{E}(\mathbf{1}_{\Omega^c}) + 2\|\Pi_m - \Pi\|_A^2. \end{aligned}$$

But $\mathbb{E}(\|\hat{\Pi}_{\hat{m}}\|^2 \mathbf{1}_{\Omega}) \leq 2\mathbb{E}(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_{\Omega}) + 2\|\Pi_m\|^2$ and $k_n = \sqrt{n}$, so

$$\begin{aligned} \mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 &\leq 2\mathbb{E} \left(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_{\Omega} \right) (1 + 2\|\Pi\|_A^2) + \frac{4\|\Pi\|_A^4}{n} \\ &\quad + 4(n + \|\Pi\|_A^2) P(\Omega^c) + 2\|\Pi_m - \Pi\|_A^2. \end{aligned}$$

We now state the following proposition :

Proposition 2. *There exists $C_0 > 0$ such that*

$$P(\Omega^c) \leq \frac{C_0}{n^2}.$$

Hence

$$(12) \quad \mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 \leq 2\|\Pi_m - \Pi\|_A^2 + 2\mathbb{E} \left(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_\Omega \right) (1 + 2\|\Pi\|_A^2) + \frac{4}{n} (\|\Pi\|_A^4 + C_0(1 + \|\Pi\|_A^2)).$$

Now we have to bound $\mathbb{E} \left(\|\hat{\Pi}_{\hat{m}} - \Pi_m\|^2 \mathbf{1}_\Omega \right)$. The estimators $\hat{\Pi}_m$ are defined by minimization of the contrast on a set Γ defined in (6). Let us prove that this set Γ includes Ω . More precisely, we prove that $\Delta \subset \Gamma$. For $T = \sum_\lambda a_\lambda \omega_\lambda \in \mathbb{S}_m$, the matrix $A_m = (a_\lambda)$ of its coefficients in the basis $(\omega_\lambda(x, y))$ verifies $\Psi_n(T) = {}^t A_m G_m A_m$. Then, on Δ ,

$${}^t A_m G_m A_m \geq \frac{2}{3} \|T\|_f^2 \geq \frac{2}{3} f_0 \|T\|^2.$$

Now, using P2, $\|T\|^2 = {}^t A_m A_m$ and then ${}^t A_m G_m A_m \geq (2/3) f_0 {}^t A_m A_m$. If μ is an eigenvalue of G_m , there exists $A_m \neq 0$ such that $G_m A_m = \mu A_m$ and then ${}^t A_m G_m A_m = \mu {}^t A_m A_m$. Then, on Δ ,

$$\mu {}^t A_m A_m \geq \frac{2}{3} f_0 {}^t A_m A_m.$$

Consequently $\mu \geq (2/3) f_0$. So $\Delta \subset \Gamma$ and $\hat{\Pi}_{\hat{m}}$ minimizes the contrast on Δ .

We now observe that, for all functions T, S

$$\gamma_n(T) - \gamma_n(S) = \Psi_n(T - S) - \frac{2}{n} \sum_{k=1}^n [V_{(T-S)}(Y_k, Y_{k+1}) - Q_{(T-S)S}(Y_k)].$$

Then, since on Δ , $\gamma_n(\hat{\Pi}_{\hat{m}}) + \text{pen}(\hat{m}) \leq \gamma_n(\Pi_m) + \text{pen}(m)$,

$$\begin{aligned} \Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m) &\leq \frac{2}{n} \sum_{k=1}^n [V_{(\hat{\Pi}_{\hat{m}} - \Pi_m)}(Y_k, Y_{k+1}) - Q_{(\hat{\Pi}_{\hat{m}} - \Pi_m)\Pi_m}(Y_k)] \\ &\quad + \text{pen}(m) - \text{pen}(\hat{m}) \\ &\leq 2Z_{n,m}(\hat{\Pi}_{\hat{m}} - \Pi_m) + \text{pen}(m) - \text{pen}(\hat{m}) \\ &\leq 2\|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f \sup_{T \in B_f(m, \hat{m})} Z_{n,m}(T) + \text{pen}(m) - \text{pen}(\hat{m}) \end{aligned}$$

where

$$Z_{n,m}(T) = \frac{1}{n} \sum_{k=1}^n [V_T(Y_k, Y_{k+1}) - Q_{T\Pi_m}(Y_k)]$$

and, for all m' , $B_f(m, m') = \{T \in S_m + S_{m'}, \quad \|T\|_f = 1\}$. Now let $p(.,.)$ be a function such that for all m, m' , $12p(m, m') \leq \text{pen}(m) + \text{pen}(m')$. Then

$$\Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m) \leq \frac{1}{3} \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f^2 + 3 \left[\sup_{T \in B_f(m, \hat{m})} Z_{n,m}^2(T) - 4p(m, \hat{m}) \right] + 2\text{pen}(m).$$

So, using the definition of $\Delta \supset \Omega$,

$$\begin{aligned} \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f^2 \mathbf{1}_\Omega &\leq \frac{3}{2} \Psi_n(\hat{\Pi}_{\hat{m}} - \Pi_m) \mathbf{1}_\Omega \\ &\leq \frac{1}{2} \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f^2 \mathbf{1}_\Omega + \frac{9}{2} \sum_{m' \in \mathcal{M}_n} \left[\sup_{T \in B_f(m, m')} Z_{n,m}^2(T) - 4p(m, m') \right] \mathbf{1}_\Omega + 3\text{pen}(m) \end{aligned}$$

Thus

$$\frac{1}{2} \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_f^2 \mathbb{1}_\Omega \leq \frac{9}{2} \sum_{m' \in \mathcal{M}_n} \left[\sup_{T \in B_f(m, m')} Z_{n,m}^2(T) - 4p(m, m') \right] \mathbb{1}_\Omega + 3\text{pen}(m)$$

And using Assumption H3,

$$(13) \quad \|\hat{\Pi}_{\hat{m}} - \Pi_m\|_A^2 \mathbb{1}_\Omega \leq 9f_0^{-1} \sum_{m' \in \mathcal{M}_n} \left[\sup_{T \in B_f(m, m')} Z_{n,m}^2(T) - 4p(m, m') \right] \mathbb{1}_\Omega + 6f_0^{-1} \text{pen}(m)$$

Now, by denoting \mathbb{E}_X the expectation conditionally to X_1, \dots, X_{n+1} , the process $Z_{n,m}(T)$ can be split in the following way :

$$Z_{n,m}(T) = Z_n^{(1)}(T) - Z_n^{(2)}(T) + Z_n^{(3)}(T) + \iint T(x, y)(\Pi - \Pi_m)(x, y)f(x)dxdy$$

with

$$\begin{cases} Z_n^{(1)}(T) = \frac{1}{n} \sum_{k=1}^n V_T(Y_k, Y_{k+1}) - \mathbb{E}_X[V_T(Y_k, Y_{k+1})] \\ Z_n^{(2)}(T) = \frac{1}{n} \sum_{k=1}^n Q_{T\Pi_m}(Y_k) - \mathbb{E}[Q_{T\Pi_m}(Y_k)] \\ Z_n^{(3)}(T) = \frac{1}{n} \sum_{k=1}^n T(X_k, X_{k+1}) - \mathbb{E}[T(X_k, X_{k+1})] \end{cases}$$

Then, by introducing functions $p_1(\cdot, \cdot)$, $p_2(\cdot, \cdot)$ and $p_3(\cdot, \cdot)$

$$\begin{aligned} \sup_{T \in B_f(m, m')} Z_{n,m}^2(T) - 4p(m, m') &\leq 4 \sup_{T \in B_f(m, m')} (Z_n^{(1)}(T)^2 - p_1(m, m')) \\ &+ 4 \sup_{T \in B_f(m, m')} (Z_n^{(2)}(T)^2 - p_2(m, m')) + 4 \sup_{T \in B_f(m, m')} (Z_n^{(3)}(T)^2 - p_3(m, m')) \\ &+ 4((p_1 + p_2 + p_3)(m, m') - p(m, m')) + 4 \sup_{T \in B_f(m, m')} \|(\Pi - \Pi_m)\mathbb{1}_A\|_f^2 \|T\|_f^2 \end{aligned}$$

We now employ the following propositions.

Proposition 3. *Let $p_1(m, m') = K_1 D_{m''}^{4\gamma+2}/n$ where $m'' = \max(m, m')$. Then, if $r > 2\gamma + 1/2$, there exists a positive constant C_1 such that*

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_n^{(1)}(T)^2 - p_1(m, m') \right]_+ \right) \leq \frac{C_1}{n}.$$

Proposition 4. *Let $p_2(m, m') = p_2^{(1)}(m, m') + p_2^{(2)}(m, m')$ with $p_2^{(1)}(m, m') = K_2 \|\Pi\|_A^2 D_{m''}^{2\gamma+7/2}/n$ and $p_2^{(2)}(m, m') = K_2 \|\Pi\|_A^2 (\sum_k \beta_k) D_{m''}^3/n$ where $m'' = \max(m, m')$. Then, if $r > 2\gamma + 3/2$, there exists a positive constant C_2 such that*

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_n^{(2)}(T)^2 - p_2(m, m') \right]_+ \mathbb{1}_\Omega \right) \leq \frac{C_2}{n}.$$

Proposition 5. *Let $p_3(m, m') = K_3 \sum_k \beta_{2k} D_{m''}^2/n$ where $m'' = \max(m, m')$. Then, there exists a positive constant C_3 such that*

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_n^{(3)}(T)^2 - p_3(m, m') \right]_+ \mathbb{1}_\Omega \right) \leq \frac{C_3}{n}.$$

The two first propositions are proved in Sections 6.6 and 6.7. The last proposition is proved in Lacour (2007b) Section 6.5 (for an other basis but only the property P3 $\|\sum_{jk} \varphi_{jk}^2\|_\infty \leq \Phi_1 D_m$ is used).

Then we get

$$\begin{aligned} \sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_{n,m}^2(T) - 4p(m, m') \right] \mathbb{1}_\Omega \right) &\leq 4 \frac{C_1 + C_2 + C_3}{n} \\ &+ 4 \|(\Pi - \Pi_m) \mathbb{1}_A\|_f^2 + 4 \sum_{m' \in \mathcal{M}_n} ((p_1 + p_2 + p_3)(m, m') - p_1(m, m')). \end{aligned}$$

But, if $\gamma > 3/4$, $4\gamma + 2 > 2\gamma + 7/2$ and there exists m_2 such that for all $m' > m_2$, $p_1(m, m') > p_2(m, m') + p_3(m, m')$. That implies that

$$\begin{aligned} &\sum_{m' \in \mathcal{M}_n} (p_1(m, m') + p_2(m, m') + p_3(m, m') - 2p_1(m, m')) \\ &\leq \sum_{m' \leq m_2} (p_2(m, m') + p_3(m, m') - p_1(m, m')) \leq \frac{C(m_2)}{n}. \end{aligned}$$

Thus in the case $\gamma > 3/4$, we choose $p = 2p_1$ and

$$(14) \quad \begin{aligned} \sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B(m')} Z_{n,m}^2(T) - 4p(m, m') \right] \mathbb{1}_\Omega \right) &\leq 4 \frac{C_1 + C_2 + C_3 + C(m_2)}{n} \\ &+ 4 \|f\|_{\infty, A_1} \|\Pi - \Pi_m\|_A^2 \end{aligned}$$

If $\gamma = 3/4$, we choose $p = 2(p_1 + p_2^{(1)})$. Since there exists m_2 such that for all $m' > m_2$, $p_1(m, m') + p_2^{(1)}(m, m') > p_2^{(2)}(m, m') + p_3(m, m')$, we can write

$$\begin{aligned} &\sum_{m' \in \mathcal{M}_n} (p_1(m, m') + p_2(m, m') + p_3(m, m') - 2p(m, m')) \\ &\leq \sum_{m' \leq m_2} (p_2^{(2)}(m, m') + p_3(m, m') - p_1(m, m') - p_2^{(1)}(m, m')) \leq \frac{C(m_2)}{n} \end{aligned}$$

and (14) holds.

Finally, combining (12), (13) and (14), we obtain

$$\begin{aligned} \mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 &\leq 2 \|\Pi_m - \Pi\|_A^2 + \frac{4}{n} (\|\Pi\|_A^4 + C_0(1 + \|\Pi\|_A^2)) \\ &+ 2(1 + 2\|\Pi\|_A^2) 9f_0^{-1} \left[4 \frac{C_1 + C_2 + C_3 + C(m_2)}{n} + 4 \|f\|_{\infty, A_1} \|\Pi - \Pi_m\|_A^2 \right] \\ &+ 2(1 + 2\|\Pi\|_A^2) 6f_0^{-1} \text{pen}(m). \end{aligned}$$

Then, by letting $C = \max(2 + 72f_0^{-1} \|f\|_{\infty, A_1} (1 + 2\|\Pi\|_A^2), 12f_0^{-1} (1 + 2\|\Pi\|_A^2))$,

$$\mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 \leq C \inf_{m \in \mathcal{M}_n} (\|\Pi_m - \Pi\|_A^2 + \text{pen}(m)) + \frac{C'}{n}$$

We still have to verify that $12p(m, m') \leq \text{pen}(m) + \text{pen}(m')$. But, if $\gamma > 3/4$,

$$12p(m, m') = 24K_1 \frac{D_{m''}^{4\gamma+2}}{n} = 24K_1 \frac{\dim(S_m + S_{m'})^{4\gamma+2}}{n} \leq \text{pen}(m) + \text{pen}(m')$$

with $\text{pen}(m) \geq 24K_1 D_m^{4\gamma+2}/n$. And if $\gamma = 3/4$,

$$12p(m, m') = 24(K_1 + K_2 \|\Pi\|_A^2) \frac{D_{m''}^5}{n} \leq \text{pen}(m) + \text{pen}(m')$$

with $\text{pen}(m) \geq 24(K_1 + K_2 \|\Pi\|_A^2) D_m^{4\gamma+2}/n$.

6.4. Proof of Corollary 1. It follows from Meyer (1990) Chapter 6, Section 10 that Π belongs to $B_{2,\infty}^\alpha$ if and only if $\sup_{j \geq J} 2^{2j\alpha} (\sum_{k,l} |a_{jkl}|^2)^{1/2} < \infty$ with $a_{jkl} = \int \Pi(x, y) \varphi_{jk}(x) \varphi_{jl}(y) dx dy$. Then

$$d_A^2(\Pi, S_m) = \sum_{j>m} \sum_{k,l} |a_{jkl}|^2 \leq C \sum_{j>m} 2^{-4j\alpha} \leq C' D_m^{-2\alpha}$$

Since $d_A^2(\Pi, S_m) = O(D_m^{-2\alpha})$, Theorem 1 becomes

$$\mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 \leq C'' \inf_{m \in \mathcal{M}_n} \{D_m^{-2\alpha} + \frac{D_m^{4\gamma+2}}{n}\}.$$

with C'' a positive constant. By setting D_{m_1} the integer part of $n^{1/(4\gamma+2\alpha+2)}$, then

$$\mathbb{E} \|\tilde{\Pi} - \Pi\|_A^2 \leq C'' \{D_{m_1}^{-2\alpha} + \frac{D_{m_1}^{4\gamma+2}}{n}\} = O(n^{-\frac{2\alpha}{4\gamma+2\alpha+2}}).$$

6.5. Proof of Proposition 2. We first remark that $P(\Omega_\rho^{*c}) \leq P(\Omega_X^{*c}) + P(\Delta^c \cap \Omega_X^*)$. In the geometric case $\beta_{q_n} \leq e^{-\theta c \ln(n)} \leq n^{-\theta c}$ and in the other case $\beta_{q_n} \leq (q_n)^{-\theta} \leq n^{-\theta c}$. Then

$$P(\Omega_X^{*c}) \leq 2p_n \beta_{q_n} \leq n^{1-c\theta}.$$

But, $c\theta > 3$ and so $P(\Omega_X^{*c}) \leq n^{-2}$. We still have to bound $P(\Delta^c \cap \Omega_X^*)$. To do this, we observe that if $\omega \in \Delta^c$, then there exists T in \mathcal{S} such that $\|T\|_f^2 > (3/2)\Psi_n(T)$ and then $\|T\|_f^2 > (3/2)\mathbb{E}_X \Psi_n(T)$. But $\mathbb{E}_X \Psi_n(T) = \frac{1}{n} \sum_{k=1}^n \int T^2(X_k, y) dy$. So $P(\Delta^c \cap \Omega_X^*) \leq P(\Delta'^c \cap \Omega_X^*)$ with

$$\Delta' = \{\forall T \in \mathcal{S} \quad \|T\|_f^2 \leq \frac{3}{2} \frac{1}{n} \sum_{k=1}^n \int T^2(X_k, y) dy\}.$$

Let us remark that $(1/n) \sum_{k=1}^n \int T^2(X_k, y) dy - \|T\|_f^2 = \nu_n(T^2)$ with

$$\nu_n(T) = \frac{1}{n} \sum_{i=1}^n \int [T(X_i, y) - \mathbb{E}(T(X_i, y))] dy.$$

Hence

$$P(\Delta'^c \cap \Omega_X^*) \leq P(\sup_{T \in \mathcal{B}} |\nu_n(T^2)| \mathbf{1}_{\Omega_X^*} > 1/3)$$

with $\mathcal{B} = \{T \in \mathcal{S} \quad \|T\|_f = 1\}$.

A function T in \mathcal{S} can be written $T(x, y) = \sum_{j=J}^{m_0} \sum_{kl} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y)$ where m_0 is such that $\mathcal{S} = \mathbb{S}_{m_0}$. Then

$$\nu_n(T^2) \mathbf{1}_{\Omega_X^*} = \sum_{jkk'l} a_{jkl} a_{jk'l} \bar{\nu}_n(\varphi_{jk} \varphi_{jk'})$$

where

$$(15) \quad \bar{\nu}_n(u) = \frac{1}{n} \sum_{i=1}^n [u(X_i^*) - \mathbb{E}(u(X_i^*))].$$

Let $b_{jk} = (\sum_l a_{jkl}^2)^{1/2}$, then $|\nu_n(T^2)| \leq \sum_{jkk'} b_{jk} b_{jk'} |\bar{\nu}_n(\varphi_{jk} \varphi_{jk'})|$ and, if $T \in \mathcal{B}$, $\sum_{jk} b_{jk}^2 = \sum_{jkl} a_{jkl}^2 = \|T\|^2 \leq f_0^{-1}$

Thus,

$$\sup_{T \in \mathcal{B}} |\nu_n(T^2)| \leq f_0^{-1} \sup_{\sum b_{jk}^2 = 1} \sum_{jkk'} b_{jk} b_{jk'} |\bar{\nu}_n(\varphi_{jk} \varphi_{jk'})|.$$

For the sake of simplicity, we denote $\lambda = (j, k)$ and $\lambda' = (j, k')$ so that

$$\sup_{T \in \mathcal{B}} |\nu_n(T^2)| \leq f_0^{-1} \sup_{\sum b_\lambda^2 = 1} \sum_{\lambda\lambda'} b_\lambda b_{\lambda'} |\bar{\nu}_n(\varphi_\lambda \varphi_{\lambda'})|.$$

Lemma 3. Let $B_{\lambda,\lambda'} = \|\varphi_\lambda \varphi_{\lambda'}\|_\infty$ and $V_{\lambda,\lambda'} = \|\varphi_\lambda \varphi_{\lambda'}\|_2$. Let, for any symmetric matrix $(A_{\lambda,\lambda'})$

$$\bar{\rho}(A) = \sup_{\sum a_\lambda^2 = 1} \sum_{\lambda,\lambda'} |a_\lambda a_{\lambda'}| A_{\lambda,\lambda'}$$

and $L(\varphi) = \max\{\bar{\rho}^2(V), \bar{\rho}(B)\}$. Then there exists $\Phi_0 > 0$ such that $L(\varphi) \leq \Phi_0 \mathcal{D}^2$.

This lemma is proved in Baraud et al. (2001) for an orthonormal basis verifying $\|\sum_\lambda \varphi_\lambda^2\|_\infty \leq \Phi_0 \mathcal{D}$, that is ensured by property P3.

Now let $x = \frac{f_0^2}{24\|f\|_{\infty,A_1}L(\varphi)}$ and

$$D = \left\{ \forall \lambda \forall \lambda' \quad |\bar{\nu}_n(\varphi_\lambda \varphi_{\lambda'})| \leq \left[B_{\lambda,\lambda'} x + V_{\lambda,\lambda'} \sqrt{2\|f\|_{\infty,A_1} x} \right] \right\}.$$

On D :

$$\begin{aligned} \sup_{T \in \mathcal{B}} |\nu_n(T^2)| &\leq f_0^{-1} \sup_{\sum b_\lambda^2 = 1} \sum_{\lambda,\lambda'} b_\lambda b_{\lambda'} \left[B_{\lambda,\lambda'} x + V_{\lambda,\lambda'} \sqrt{2\|f\|_{\infty,A_1} x} \right] \\ &\leq f_0^{-1} \left[\bar{\rho}(B)x + \bar{\rho}(V) \sqrt{2\|f\|_{\infty,A_1} x} \right] \\ &\leq \frac{f_0}{24\|f\|_{\infty,A_1}L(\varphi)} \frac{\bar{\rho}(B)}{L(\varphi)} + \frac{1}{\sqrt{12}} \left(\frac{\bar{\rho}^2(V)}{L(\varphi)} \right)^{1/2} \leq \frac{1}{24} + \frac{1}{2\sqrt{3}} < \frac{1}{3}. \end{aligned}$$

Then $P\left(\sup_{T \in \mathcal{B}} |\nu_n(T^2)| > 1/3\right) \leq P(D^c)$. But $\bar{\nu}_n(u) = \bar{\nu}_{n,1}(u)/2 + \bar{\nu}_{n,2}(u)/2$ with

$$\bar{\nu}_{n,s}(u) = \frac{1}{p_n} \sum_{l=0}^{p_n-1} Y_{l,s}(u) \quad s = 1, 2$$

$$\text{with } \begin{cases} Y_{l,1}(u) &= \frac{1}{q_n} \sum_{i=2lq_n+1}^{(2l+1)q_n} [u(X_i^*) - \mathbb{E}(u(X_i^*))], \\ Y_{l,2}(u) &= \frac{1}{q_n} \sum_{i=2(2l+1)q_n+1}^{(2l+2)q_n} [u(X_i^*) - \mathbb{E}(u(X_i^*))]. \end{cases}$$

To bound $P(\bar{\nu}_{n,1}(\varphi_\lambda \varphi_{\lambda'}) \geq B_{\lambda,\lambda'} x + V_{\lambda,\lambda'} \sqrt{2\|f\|_{\infty,A_1} x})$, we will use the Bernstein inequality given in Birgé and Massart (1998). A fast computation gives $\mathbb{E}|Y_{l,1}(\varphi_\lambda \varphi_{\lambda'})|^p \leq 2^{p-2} (B_{\lambda,\lambda'})^{p-2} (\sqrt{\|f\|_{\infty,A_1}} V_{\lambda,\lambda'})^2$. And then

$$P(|\bar{\nu}_{n,s}(\varphi_\lambda \varphi_{\lambda'})| \geq B_{\lambda,\lambda'} x + V_{\lambda,\lambda'} \sqrt{2\|f\|_{\infty,A_1} x}) \leq 2e^{-p_n x}.$$

Let $C = f_0^2 [48\|f\|_{\infty,A_1}]^{-1}$, so that $x = 2C/L(\varphi)$. Given that $P(\Delta^c \cap \Omega_X^*) \leq P(D^c) \leq \sum_{\lambda,\lambda'} P\left(|\bar{\nu}_n(\varphi_\lambda \varphi_{\lambda'})| > B_{\lambda,\lambda'} x + V_{\lambda,\lambda'} \sqrt{2\|f\|_{\infty,A_1} x}\right)$

$$P(\Delta^c \cap \Omega_X^*) \leq 4\mathcal{D}^2 \exp\left\{-\frac{2p_n C}{L(\varphi)}\right\} \leq 4n^{1/(2\gamma+1)} \exp\left\{-C \frac{n}{q_n L(\varphi)}\right\}.$$

But $L(\varphi) \leq \Phi_0 \mathcal{D}^2 \leq \Phi_0 n^{1/(2\gamma+1)}$ and $q_n \leq n^{1/2}$ so

$$P(\Delta^c \cap \Omega_X^*) \leq 4n^{1/(2\gamma+1)} \exp\left\{-\frac{C}{\Phi_0} n^{\frac{2\gamma-1}{2(2\gamma+1)}}\right\} \leq \frac{C'}{n^2}$$

because $\gamma > 1/2$.

6.6. Proof of Proposition 3. First we need to isolate even terms from odd terms in $Z_n^{(1)}(T)$ to avoid overlaps: $Z_n^{(1)}(T)\mathbf{1}_{\Omega^*} = \frac{1}{2}(Z_n^{(1,1)}(T) + Z_n^{(1,2)}(T))$ with

$$\begin{cases} Z_n^{(1,1)}(T) = \frac{1}{n} \sum_{i=1, i \text{ odd}}^n V_T(Y_i, Y_{i+1}) - \mathbb{E}_X[V_T(Y_i, Y_{i+1})] \\ Z_n^{(1,2)}(T) = \frac{1}{n} \sum_{i=1, i \text{ even}}^n V_T(Y_i, Y_{i+1}) - \mathbb{E}_X[V_T(Y_i, Y_{i+1})] \end{cases}$$

It is sufficient to deal with the first term only, as the second one is similar. For each i , let $U_i = (Y_{2i-1}, Y_{2i})$, then

$$Z_n^{(1,1)}(T) = \frac{1}{n/2} \sum_{i=1}^{n/2} \{V_T(U_i) - \mathbb{E}_X[V_T(U_i)]\}.$$

Notice that conditionally to X_1, \dots, X_n , the U_i 's are independent. Thus we can use the Talagrand inequality recalled in Lemma 5 to bound

$$\mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_n^{(1,1)}(T)^2 - p_1(m, m') \right]_+ \right).$$

We first remark that Property P1 entails $B_f(m, m') \subset \mathbb{S}_{m''}$ with $m'' = \max(m, m')$. Then, if T belongs to $B_f(m, m')$,

$$T(x, y) = \sum_{j=J}^{m''} \sum_{kl} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y)$$

with $\sum_{jkl} a_{jkl}^2 = \|T\|^2 \leq f_0^{-1}$.

- Let us bound $\|V_T\|_\infty$ for T in $B_f(m, m')$. If $T(x, y) = \sum_{jkl} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y)$,

$$|V_T(x, y)|^2 \leq \sum_{jkl} a_{jkl}^2 \sum_{jkl} |V_{\varphi_{jk} \otimes \varphi_{jl}}(x, y)|^2.$$

Then, since $V_{s \otimes t}(x, y) = v_s(x) v_t(y)$,

$$\sup_{T \in B_f(m, m')} |V_T(x, y)|^2 \leq f_0^{-1} \sum_{jkl} |v_{\varphi_{jk}}(x) v_{\varphi_{jl}}(y)|^2.$$

But, according to Property P4, $\|\sum_k |v_{\varphi_{jk}}|^2\|_\infty \leq \Phi_1(2^j)^{2\gamma+2}$. So

$$\sup_{T \in B_f(m, m')} \|V_T\|_\infty^2 \leq f_0^{-1} \Phi_1^2 \sum_{j=J}^{m''} (2^j)^{4\gamma+4} \leq f_0^{-1} \Phi_1^2 \frac{2^{4\gamma+4}}{2^{4\gamma+4} - 1} D_{m''}^{4\gamma+4}$$

and $M_1 = f_0^{-1/2} \Phi_1 \sqrt{2^{4\gamma+4}/(2^{4\gamma+4} - 1)} D_{m''}^{2\gamma+2}$.

- To compute H^2 we write

$$\begin{aligned} \mathbb{E}_X \left[\sup_{T \in B_f(m, m')} Z_n^{(1,1)}(T)^2 \right] &\leq f_0^{-1} \sum_{jkl} \mathbb{E}_X [Z_n^{(1,1)}(\varphi_{jk} \otimes \varphi_{jl})^2] \\ &\leq f_0^{-1} \sum_{jkl} \text{Var}_X \left(\frac{1}{n} \sum_{i=1, i \text{ odd}}^n v_{\varphi_{jk}}(Y_i) v_{\varphi_{jl}}(Y_{i+1}) \right) \\ (16) \quad &\leq f_0^{-1} \sum_{jkl} \frac{1}{n} \text{Var}_X (v_{\varphi_{jk}}(Y_1) v_{\varphi_{jl}}(Y_2)) \leq \frac{f_0^{-1}}{n} \sum_{jkl} \mathbb{E}_X [|v_{\varphi_{jk}}(Y_1)|^2 |v_{\varphi_{jl}}(Y_2)|^2] \end{aligned}$$

Here Var_X denotes the variance conditionally to X_1, \dots, X_{n+1} . Now, for any function G , the following relation holds

$$\begin{aligned} \mathbb{E}_X[|G|^2(Y_1, Y_2)] &= \mathbb{E}_X[|G|^2(X_1 + \varepsilon_1, X_2 + \varepsilon_2)] \\ &= \iint |G|^2(X_1 + z_1, X_2 + z_2) q(z_1) q(z_2) dz_1 dz_2 \\ &= \iint |G|^2(u_1, u_2) q(u_1 - X_1) q(u_2 - X_2) du_1 du_2 \leq \|q\|_\infty^2 \|G\|^2 \end{aligned}$$

Then, coming back to (16),

$$\begin{aligned} \mathbb{E}_X \left[\sup_{T \in B_f(m, m')} Z_n^{(1,1)}(T)^2 \right] &\leq \frac{f_0^{-1}}{n} \|q\|_\infty^2 \sum_{jkl} \|v_{\varphi_{jk}} \otimes v_{\varphi_{jl}}\|^2 \\ &\leq \frac{f_0^{-1} \|q\|_\infty^2}{n} \sum_j \left(\sum_k \|v_{\varphi_{jk}}\|^2 \right)^2 \leq \frac{\Phi_1^2 f_0^{-1} \|q\|_\infty^2}{n} \sum_{j=J}^{m''} (2^j)^{4\gamma+2}, \end{aligned}$$

using P5. Then $H^2 = \Phi_1^2 f_0^{-1} \|q\|_\infty^2 2^{4\gamma+2} / (2^{4\gamma+2} - 1) \frac{D_{m''}^{4\gamma+2}}{n}$.

- We still have to find v . First

$$\text{Var}_X(V_T(Y_i, Y_{i+1})) \leq \mathbb{E}_X |V_T(Y_i, Y_{i+1})|^2 \leq \|q\|_\infty^2 \|V_T\|^2$$

We now observe that $\|V_T\|^2 = \|V_T^*\|^2 / (4\pi^2)$ and then

$$\begin{aligned} \|V_T\|^2 &= \frac{1}{4\pi^2} \iint \left| \frac{T^*(u, v)}{q^*(-u)q^*(-v)} \right|^2 dudv \\ &\leq \frac{1}{4\pi^2} \sqrt{\iint \frac{|T^*(u, v)|^2}{|q^*(-u)q^*(-v)|^4} dudv} \sqrt{\iint |T^*(u, v)|^2 dudv} \\ &\leq \frac{1}{4\pi^2} \sqrt{\sum_{jkl} a_{jkl}^2 \sum_{jkl} \iint \frac{|\varphi_{jk}^*(u)\varphi_{jl}^*(v)|^2}{|q^*(-u)q^*(-v)|^4} dudv} \sqrt{4\pi^2 \|T\|^2} \end{aligned}$$

For $T \in B_f(m, m')$,

$$\|V_T\|^2 \leq \frac{f_0^{-1/2}}{2\pi} \sqrt{f_0^{-1} \sum_j \sum_{kl} \int \frac{|\varphi_{jk}^*(u)|^2}{|q^*(-u)|^4} du \int \frac{|\varphi_{jl}^*(u)|^2}{|q^*(-u)|^4} du}.$$

But $(\varphi_{jk})^*(u) = 2^{-j/2} e^{iuk/2^j} \varphi^*(u/2^j)$ and then

$$\begin{aligned} \int \frac{|\varphi_{jk}^*(u)|^2}{|q^*(-u)|^4} du &\leq \int \frac{2^{-j} |\varphi^*(u/2^j)|^2}{|q^*(-u)|^4} du \\ &\leq \int \frac{|\varphi^*(v)|^2}{|q^*(-v2^j)|^4} dv \leq k_0^{-4} (2^j)^{4\gamma} \int |\varphi^*(v)|^2 (v^2 + 1)^{2\gamma} dv \end{aligned}$$

Since $r > 2\gamma + 1/2$, Lemma 4 gives

$$\sum_{(k,l) \in \Lambda_j} \int \frac{|\varphi_{jk}^*(u)|^2}{|q^*(-u)|^4} du \int \frac{|\varphi_{jl}^*(u)|^2}{|q^*(-u)|^4} du \leq 3 \cdot 2^{2j} C_{2,4\gamma}^2 k_0^{-8} (2^j)^{8\gamma}$$

Then

$$\|V_T\|^2 \leq \frac{C_{2,4\gamma} f_0^{-1} k_0^{-4}}{2\pi} \sqrt{\sum_{j=J}^{m''} 3(2^j)^{8\gamma+2}} \leq \frac{C_{2,4\gamma} f_0^{-1} k_0^{-4}}{2\pi} \left(\frac{3 \cdot 2^{8\gamma+2}}{2^{8\gamma+2} - 1} \right)^{1/2} D_{m''}^{4\gamma+1}$$

and $v = \|q\|_\infty^2 C_{2,4\gamma} f_0^{-1} k_0^{-4} \sqrt{3 \cdot 2^{8\gamma+2}} D_{m''}^{4\gamma+1} / (2\pi \sqrt{2^{8\gamma+2} - 1})$.

We can now apply inequality (19)

$$\begin{aligned} \mathbb{E} \left[\sup_{T \in B_f(m, m')} |Z_n^{(1,1)}(T)|^2 - 6H^2 \right]_+ &\leq C \left(\frac{v}{n} e^{-k_1 \frac{nH^2}{v}} + \frac{M_1^2}{n^2} e^{-k_2 \frac{nH}{M_1}} \right) \\ &\leq C' \left(\frac{D_{m''}^{4\gamma+1}}{n} e^{-k'_1 D_{m''}} + \frac{D_{m''}^{4\gamma+4}}{n^2} e^{-k'_2 \sqrt{n}/D_{m''}} \right). \end{aligned}$$

But there exists a positive constant K such that

$$\sum_{m' \in \mathcal{M}_n} D_{m''}^{4\gamma+1} e^{-k'_1 D_{m''}} \leq K.$$

Moreover, since $D_{m''} \leq n^{\frac{1}{4\gamma+2}}$, $D_{m''}^{4\gamma+4} e^{-k'_2 \sqrt{n}/D_{m''}} / n \leq n^{1/(2\gamma+1)} e^{-k'_2 n^{\gamma/(2\gamma+1)}}$ so that

$$\sum_{m' \in \mathcal{M}_n} D_{m''}^{4\gamma+4} e^{-k'_2 \sqrt{n}/D_{m''}} / n^2 \leq K' / n.$$

Then, setting $K_1 = 6\Phi_1^2 f_0^{-1} \|q\|_\infty^2 2^{4\gamma+2} / (2^{4\gamma+2} - 1)$,

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left[\sup_{T \in B_f(m, m')} |Z_n^{(1,1)}(T)|^2 - K_1 \frac{D_{m''}^{4\gamma+2}}{n} \right]_+ \leq \frac{C''}{n}$$

and the proposition is proved.

6.7. Proof of Proposition 4. Since Π_m belongs to \mathbb{S}_m , it can be written

$$\Pi_m(x, y) = \sum_{j'=J}^m \sum_{(k', l') \in \Lambda_{j'}} b_{j'k'l'} \varphi_{j'k'}(x) \varphi_{j'l'}(y)$$

with $\sum_{j'k'l'} b_{j'k'l'}^2 = \|\Pi_m\|^2 \leq \|\Pi\|_A^2$. From the embedding $B_f(m, m') \subset S_{m''}$ (where $m'' = \max(m, m')$), we have, if T belongs to $B_f(m, m')$,

$$T(x, y) = \sum_{j=J}^{m''} \sum_{(k, l) \in \Lambda_j} a_{jkl} \varphi_{jk}(x) \varphi_{jl}(y)$$

with $\sum_{jkl} a_{jkl}^2 = \|T\|^2 \leq f_0^{-1}$.

We use the Talagrand inequality (19) in Lemma 5. But the variables Y_i are not independent. We shall use the following approximation variables

$$\forall 1 \leq i \leq n+1 \quad Y_i^* = X_i^* + \varepsilon_i.$$

These variables have the same properties (see (11)) that the X_i^* s. More precisely, let, for $l = 0, \dots, p_n - 1$, $C_l = (Y_{2lq_n+1}, \dots, Y_{(2l+1)q_n})$, $D_l = (Y_{(2l+1)q_n+1}, \dots, Y_{(2l+2)q_n})$, $C_l^* = (Y_{2lq_n+1}^*, \dots, Y_{(2l+1)q_n}^*)$, $D_l^* = (Y_{(2l+1)q_n+1}^*, \dots, Y_{(2l+2)q_n}^*)$. Then, since A_l and A_l^* have the same distribution and the sequences (ε_i) and (X_i) are independent, C_l and C_l^* have the same distribution. Moreover the construction of A_l^* via Berbee's coupling Lemma implies that C_l^* and $C_{l'}^*$ are independent if $l \neq l'$.

Now we split $Z_n^{(2)}$ into two terms: $Z_n^{(2)}(T) \mathbb{1}_\Omega = (1/2) Z_n^{(2,1)}(T) + (1/2) Z_n^{(2,2)}(T)$ where

$$\begin{cases} Z_n^{(2,1)}(T) = \frac{1}{p_n} \sum_{l=0}^{p_n-1} \frac{1}{q_n} \sum_{i=2lq_n+1}^{(2l+1)q_n} Q_{T\Pi_m}(Y_i^*) - \mathbb{E}[Q_{T\Pi_m}(Y_i^*)] \\ Z_n^{(2,2)}(T) = \frac{1}{p_n} \sum_{l=0}^{p_n-1} \frac{1}{q_n} \sum_{i=(2l+1)q_n+1}^{(2l+2)q_n} Q_{T\Pi_m}(Y_i^*) - \mathbb{E}[Q_{T\Pi_m}(Y_i^*)] \end{cases}$$

Then we apply Talagrand's inequality to $Z_n^{(2,1)}(T)$.

- Let us first compute M_1 . We have to bound $\|Q_{T\Pi_m}\|_\infty$ for T in $B_f(m, m')$. By linearity of Q

$$Q_{T\Pi_m}(x) = \sum_{jkl} a_{jkl} \sum_{j'k'l'} b_{j'k'l'} Q_{\varphi_{jk}\varphi_{j'k'} \otimes \varphi_{jl}\varphi_{j'l'}}(x)$$

Then, since $Q_{s \otimes t}(x) = v_s(x) \int t(y) dy$, using the Schwarz inequality,

$$\begin{aligned} |Q_{T\Pi_m}(x)|^2 &\leq \sum_{jkl} a_{jkl}^2 b_{j'k'l'}^2 \sum |v_{\varphi_{jk}\varphi_{j'k'}}(x) \int \varphi_{jl}\varphi_{j'l'}|^2 \\ &\leq f_0^{-1} \|\Pi\|_A^2 \sum_{jkk'l} |v_{\varphi_{jk}\varphi_{j'k'}}(x)|^2 \end{aligned}$$

since the φ_{jl} are orthonormal. The property P6 gives then

$$\|Q_{T\Pi_m}\|_\infty^2 \leq f_0^{-1} \|\Pi\|_A^2 \Phi_1 \sum_{j=J}^{m''} (2^j)^{2\gamma+3} 2^j$$

so that $M_1 = f_0^{-1/2} \|\Pi\|_A \sqrt{\Phi_1 2^{2\gamma+4} / (2^{2\gamma+4} - 1)} D_{m''}^{\gamma+2}$.

- Now, we compute H^2 . For $T \in B_f(m, m')$,

$$|Z_n^{(2,1)}(T)|^2 \leq \sum_{jkl} a_{jkl}^2 \sum_{jkl} |Z_n^{(2,1)}(\varphi_{jk} \otimes \varphi_{jl})|^2$$

Thus

$$\begin{aligned} \mathbb{E} \left[\sup_{T \in B_f(m, m')} Z_n^{(2,1)}(T)^2 \right] &\leq f_0^{-1} \sum_{jkl} \mathbb{E} \left[Z_n^{(2,1)}(\varphi_{jk} \otimes \varphi_{jl})^2 \right] \\ &\leq f_0^{-1} \sum_{jkl} \text{Var} \left[\frac{1}{p_n} \sum_{l=0}^{p_n-1} \frac{1}{q_n} \sum_{i=2lq_n+1}^{(2l+1)q_n} Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i^*) \right] \end{aligned}$$

The variables (C_l^*) are independent and identically distributed so

$$\mathbb{E} \left[\sup_{T \in B_f(m, m')} Z_n^{(2,1)}(T)^2 \right] \leq f_0^{-1} \sum_{jkl} \frac{1}{p_n} \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i^*) \right]$$

But, on Ω , C_1 and C_1^* have the same distribution, so that

$$\text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i^*) \right] = \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i) \right].$$

But, coming back to the definition of Q_T , for $i_1 \neq i_2$,

$$\begin{aligned} &\text{Cov}(Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_{i_1}), Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_{i_2})) \\ &= \frac{1}{4\pi^2} \iint \mathbb{E}(e^{iY_{i_1}u} e^{-iY_{i_2}v}) \frac{[(\varphi_{jk} \otimes \varphi_{jl})\Pi_m]^*(u, 0)}{q^*(-u)} \frac{[(\varphi_{jk} \otimes \varphi_{jl})\Pi_m]^*(-v, 0)}{q^*(v)} dudv \\ &= \frac{1}{4\pi^2} \iint \mathbb{E}(e^{iX_{i_1}u} e^{-iX_{i_2}v}) [(\varphi_{jk} \otimes \varphi_{jl})\Pi_m]^*(u, 0) [(\varphi_{jk} \otimes \varphi_{jl})\Pi_m]^*(-v, 0) dudv \end{aligned}$$

since $\mathbb{E}(e^{i\varepsilon_{i_1}u} e^{-i\varepsilon_{i_2}v}) = q^*(-u)q^*(v)$. Now using (10),

$$\begin{aligned} &\text{cov}(Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_{i_1}), Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_{i_2})) \\ &= \text{cov}\left(\int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_{i_1}, y) dy, \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_{i_2}, y) dy\right) \end{aligned}$$

It implies that

$$\begin{aligned} \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i) \right] &\leq \frac{1}{q_n^2} \sum_{i=1}^{q_n} \text{Var}[Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_i)] \\ &\quad + \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right] \end{aligned}$$

And then

$$\begin{aligned} (17) \quad \mathbb{E} \left[\sup_{T \in B_f(m, m')} Z_n^{(2,1)}(T)^2 \right] &\leq f_0^{-1} \sum_{jkl} \frac{1}{p_n q_n} \text{Var}[Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_1)] \\ &\quad + f_0^{-1} \sum_{jkl} \frac{1}{p_n} \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right] \end{aligned}$$

For the second term in (17), we use Lemma 6 to write

$$\begin{aligned} &\sum_{jkl} \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right] \\ &\leq \frac{4 \sum_k \beta_k}{q_n} \left\| \sum_{jkl} \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right\|_\infty \end{aligned}$$

But

$$\int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy = \sum_{k'} b_{jk'l} \varphi_{jk} \varphi_{jk'}(X_i)$$

so that

$$\sum_{jkl} \left| \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right|^2 \leq \|\Pi\|_A^2 \sum_{jkk'l} |\varphi_{jk} \varphi_{jk'}(X_i)|^2 \leq \|\Pi\|_A^2 \Phi_1^2 D_{m''}^3$$

using property P3. Therefore

$$\sum_{jkl} \text{Var} \left[\frac{1}{q_n} \sum_{i=1}^{q_n} \int (\varphi_{jk} \otimes \varphi_{jl})\Pi_m(X_i, y) dy \right] \leq \frac{4 \sum_k \beta_k}{q_n} \|\Pi\|_A^2 \Phi_1^2 D_{m''}^3$$

Then we have bound the second term in (17) by $2f_0^{-1} \sum_k \beta_k \|\Pi\|_A^2 \Phi_1^2 D_{m''}^3 / n$.

For the first term in (17), we bound $\sum_{jkl} \mathbb{E}[|Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_1)|^2]$:

$$\begin{aligned} \sum_{jkl} \mathbb{E}[|Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_1)|^2] &\leq \sum_{j'k'l'} b_{j'k'l'}^2 \sum_{jkl} \sum_{j'k'l'} \left| \int \varphi_{jl} \varphi_{j'l'} \right|^2 \mathbb{E}[|v_{\varphi_{jk} \varphi_{j'k'}}(Y_1)|^2] \\ &\leq \|\Pi\|_A^2 \sum_{jkk'} 2^j \mathbb{E}[|v_{\varphi_{jk} \varphi_{jk'}}(Y_1)|^2] \end{aligned}$$

But $\mathbb{E}[|v_{\varphi_{jk} \varphi_{jk'}}(Y_1)|^2] = \int |v_{\varphi_{jk} \varphi_{jk'}}(x)|^2 p(x) dx$ where p is the density of Y_1 . Since $p = q * f$, $|p(x)| \leq \|q\|_\infty$ for all x . Then

$$\mathbb{E}[|v_{\varphi_{jk} \varphi_{jk'}}(Y_1)|^2] \leq \|q\|_\infty \int |v_{\varphi_{jk} \varphi_{jk'}}(x)|^2 dx$$

and

$$\begin{aligned} \sum_{jkl} \mathbb{E}[|Q_{(\varphi_{jk} \otimes \varphi_{jl})\Pi_m}(Y_1)|^2] &\leq \|\Pi\|_A^2 \|q\|_\infty \sum_{j=J}^{m''} 2^j \sum_{kk'} \int |v_{\varphi_{jk} \varphi_{jk'}}(x)|^2 dx \\ &\leq \|\Pi\|_A^2 \|q\|_\infty \Phi_1 \sum_{j=J}^{m''} (2^j)^{2\gamma+3} \leq \|\Pi\|_A^2 \|q\|_\infty \Phi_1 \frac{2^{2\gamma+3}}{2^{2\gamma+3}-1} D_{m''}^{2\gamma+3}, \end{aligned}$$

applying Property P7. We finally obtain

$$\mathbb{E} \left[\sup_{T \in B_f(m, m')} Z_n^{(2,1)}(T)^2 \right] \leq 2f_0^{-1} \|\Pi\|_A^2 \|q\|_\infty \Phi_1 \frac{2^{2\gamma+3}}{2^{2\gamma+3}-1} \frac{D_{m''}^{2\gamma+3}}{n} + 2f_0^{-1} \sum_k \beta_k \|\Pi\|_A^2 \Phi_1^2 \frac{D_{m''}^3}{n}$$

Since the order of nH^2 has to be larger than the one of v , we choose

$$H^2 = 2f_0^{-1} \|\Pi\|_A^2 \Phi_1 \max(\|q\|_\infty 2^{2\gamma+3}/(2^{2\gamma+3}-1), \Phi_1) \left[\frac{D_{m''}^{2\gamma+7/2}}{n} + \left(\sum_k \beta_k \right) \frac{D_{m''}^3}{n} \right].$$

- Lastly, using Lemma 6 again,

$$\begin{aligned} \text{Var}\left(\frac{1}{q_n} \sum_{i=2lq_n+1}^{(2l+1)q_n} Q_{T\Pi_m}^2(Y_i^*)\right) &= \text{Var}\left(\frac{1}{q_n} \sum_{i=2lq_n+1}^{(2l+1)q_n} Q_{T\Pi_m}^2(Y_i)\right) \\ &\leq \frac{4}{q_n} \mathbb{E}[|Q_{T\Pi_m}|^2(Y_1)b(Y_1)] \leq \frac{4}{q_n} \|Q_{T\Pi_m}\|_\infty (\mathbb{E}[|Q_{T\Pi_m}|^2(Y_1)])^{1/2} (\mathbb{E}[b^2(Y_1)])^{1/2} \\ (18) \quad &\leq \frac{4\sqrt{2\sum_k(k+1)\beta_k}}{q_n} \|Q_{T\Pi_m}\|_\infty (\mathbb{E}[|Q_{T\Pi_m}|^2(Y_1)])^{1/2} \end{aligned}$$

We have already proved that $\|Q_{T\Pi_m}\|_\infty \leq f_0^{-1/2} \|\Pi\|_A \sqrt{\Phi_1 2^{2\gamma+4}/(2^{2\gamma+4}-1)} D_{m''}^{\gamma+2}$.

Now we need a sharp bound on $\mathbb{E}[|Q_{T\Pi_m}(Y_1)|^2]$. We have

$$\mathbb{E}[|Q_{T\Pi_m}(Y_1)|^2] \leq \|q\|_\infty \int |Q_{T\Pi_m}|^2 = \frac{\|q\|_\infty}{2\pi} \int \left| \frac{(T\Pi_m)^*(u, 0)}{q^*(-u)} \right|^2 du$$

Then it follows from the Schwarz inequality that

$$\mathbb{E}[|Q_{T\Pi_m}(Y_1)|^2] \leq \frac{\|q\|_\infty}{2\pi} \sqrt{\int \frac{|(T\Pi_m)^*(u, 0)|^2}{|q^*(-u)|^4} du} \sqrt{\int |(T\Pi_m)^*(u, 0)|^2 du}$$

We will evaluate the two terms under the square roots. First observe that

$$\begin{aligned} (T\Pi_m)^*(u, 0) &= \sum_{jkl} \sum_{j'k'l'} a_{jkl} b_{j'k'l'} (\varphi_{jk} \varphi_{j'k'})^*(u) (\varphi_{jl} \varphi_{j'l'})^*(0) \\ &= \sum_{jkk'l} a_{jkl} b_{jkk'l} (\varphi_{jk} \varphi_{jk'})^*(u) \end{aligned}$$

since $(\varphi_{jl} \varphi_{j'l'})^*(0) = \int \varphi_{jl} \varphi_{j'l'} = \mathbf{1}_{j=j', l=l'}$. Then

$$\begin{aligned} \int |(T\Pi_m)^*(u, 0)|^2 du &\leq \sum_{jkk'l} a_{jkl}^2 b_{jkk'l}^2 \sum_{jkk'l} \int |(\varphi_{jk} \varphi_{jk'})^*(u)|^2 du \\ &\leq 2\pi f_0^{-1} \|\Pi\|_A^2 \sum_{jkk'l} \int |(\varphi_{jk} \varphi_{jk'})(u)|^2 du \\ &\leq 2\pi f_0^{-1} \|\Pi\|_A^2 \sum_j 2^j \sum_{k'} \|\varphi_{jk'}\|_\infty^2 \sum_{k=k'-2N+2}^{k'+2N-2} \int |\varphi_{jk'}|^2 \end{aligned}$$

by taking account of the superposition of the supports. Using now (8)

$$\begin{aligned} \int |(T\Pi_m)^*(u, 0)|^2 du &\leq 2\pi f_0^{-1} \|\Pi\|_A^2 \sum_j 2^j C'(\varphi) 2^j (4N-3) \\ &\leq 2\pi f_0^{-1} \|\Pi\|_A^2 \Phi_1(4N-3) \frac{2}{3} D_{m''}^2 \end{aligned}$$

Now

$$\begin{aligned} \int \frac{|(T\Pi_m)^*(u, 0)|^2}{|q^*(-u)|^4} du &\leq \sum_{jkk'l} a_{jkl}^2 b_{jk'l}^2 \sum_{jkk'l} \int \frac{|(\varphi_{jk}\varphi_{jk'})^*(u)|^2}{|q^*(-u)|^4} du \\ &\leq f_0^{-1} \|\Pi\|_A^2 \sum_{jkk'l} \int 2^j \frac{|(\varphi_{jk}\varphi_{jk'})^*(2^j v)|^2}{|q^*(-2^j v)|^4} dv \end{aligned}$$

Hence, inequality (9) and Assumption H1 show that

$$\begin{aligned} &\int \frac{|(T\Pi_m)^*(u, 0)|^2}{|q^*(-u)|^4} du \\ &\leq f_0^{-1} \|\Pi\|_A^2 \sum_{jkk'l} \int 2^j C_r^2 \left[|v|^{2(1-r)} \mathbf{1}_{|v|>1} + \mathbf{1}_{|v|\leq 1} \right] k_0^{-4} ((2^j v)^2 + 1)^{2\gamma} dv \\ &\leq f_0^{-1} \|\Pi\|_A^2 C_r^2 k_0^{-4} C \sum_{jkk'l} (2^j)^{4\gamma+1} \end{aligned}$$

with $C = \int [|v|^{2(1-r)} \mathbf{1}_{|v|>1} + \mathbf{1}_{|v|\leq 1}] (v^2 + 1)^{2\gamma} dv < \infty$ as soon as $r > 2\gamma + 3/2$. Then

$$\begin{aligned} \int \frac{|(T\Pi_m)^*(u, 0)|^2}{|q^*(-u)|^4} du &\leq f_0^{-1} \|\Pi\|_A^2 C_r^2 k_0^{-4} C \sum_{j=J}^m \sum_{k'l} \sum_{k=k'-2N+2}^{k'+2N-2} (2^j)^{4\gamma+1} \\ &\leq f_0^{-1} \|\Pi\|_A^2 C_r^2 k_0^{-4} C 3(4N-3) \frac{2^{4\gamma+3}}{2^{4\gamma+3}-1} D_{m''}^{4\gamma+3} \end{aligned}$$

Finally

$$\mathbb{E}[|Q_{T\Pi_m}(Y_1)|^2] \leq \frac{\|q\|_\infty}{\pi} f_0^{-1} \|\Pi\|_A^2 (4N-3) C_r k_0^{-2} \sqrt{C' \frac{2^{4\gamma+3}}{2^{4\gamma+3}-1} \pi \Phi_1 D_{m''}^{2\gamma+5/2}}$$

Then (18) gives

$$v = \sqrt{2 \sum_k (k+1) \beta_k \|q\|_\infty f_0^{-1} \|\Pi\|_A^2 k_0^{-1} C(\gamma, r, N, \Phi_1) \frac{D_{m''}^{2\gamma+13/4}}{q_n}}.$$

Then replacing n by p_n in inequality (19) gives

$$\begin{aligned} \mathbb{E}[\sup_{T \in B_f(m, m')} |Z_n^{(2,1)}(T)|^2 - 6H^2]_+ &\leq C \left(\frac{v}{p_n} e^{-k_1 \frac{p_n H^2}{v}} + \frac{M_1^2}{p_n^2} e^{-k_2 \frac{p_n H}{M_1}} \right) \\ &\leq C' \left(\frac{D_{m''}^{2\gamma+13/4}}{n} e^{-k'_1 D_{m''}^{1/4}} + \frac{D_{m''}^{2\gamma+4} q_n^2}{n^2} e^{-k'_2 \frac{\sqrt{n}}{q_n D_{m''}^{1/4}}} \right) \end{aligned}$$

where C' and k'_1 depend on $r, N, \gamma, \Phi_1, f_0, \|\Pi\|_A, \|q\|_\infty, \sum_k (k+1) \beta_k$ and $\sum_k \beta_k$. But there exists a positive constant K such that

$$\sum_{m' \in \mathcal{M}_n} D_{m''}^{2\gamma+13/4} e^{-k'_1 D_{m''}^{1/4}} \leq K.$$

Moreover $D_{m''}^{1/4} \leq n^{1/8}$ and $q_n \leq n^c$ with $c + 1/8 < 1/2$, which involves

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left[\sup_{T \in B_f(m, m')} |Z_n^{(2,1)}(T)|^2 - K_2 \|\Pi\|_A^2 \left(\frac{D_{m''}^{2\gamma+7/2}}{n} + \left(\sum_k \beta_k \right) \frac{D_{m''}^3}{n} \right) \right]_+ \leq \frac{C''}{n}$$

with $K_2 = 12f_0^{-1}\Phi_1 \max(\|q\|_\infty 2^{2\gamma+3}/(2^{2\gamma+3}-1), \Phi_1)$. Thus, if $p_2(m, m') = p_2^{(1)}(m, m') + p_2^{(2)}(m, m')$ with $p_2^{(1)}(m, m') = K_2 \|\Pi\|_A^2 D_{m''}^{2\gamma+7/2}/n$ and $p_2^{(2)}(m, m') = K_2 \|\Pi\|_A^2 (\sum_k \beta_k) D_{m''}^3/n$, then

$$\sum_{m' \in \mathcal{M}_n} \mathbb{E} \left(\left[\sup_{T \in B_f(m, m')} Z_n^{(2)}(T)^2 - p_2(m, m') \right]_+ \mathbf{1}_\Omega \right) \leq \frac{C_2}{n}.$$

6.8. Technical Lemmas.

Lemma 4. *If $|\varphi^*(x)| \leq k_3(x^2 + 1)^{-r/2}$ for all real x then*

- if s and α are reals such that $sr > \alpha + 1$

$$\int |\varphi^*(x)|^s (x^2 + 1)^{\alpha/2} dx \leq C_{s, \alpha} < \infty$$

- if $r > 1$

$$\int |\varphi^*(y)\varphi^*(x-y)| dy \leq C_r(|x|^{1-r} \mathbf{1}_{|x|>1} + \mathbf{1}_{|x|\leq 1})$$

Proof of Lemma 4:

• For the first point, it is sufficient to observe that the function $(x^2 + 1)^{(-rs+\alpha)/2}$ is integrable if $-rs + \alpha > -1$.

- By changing the variable ($y = xu$), we get

$$\begin{aligned} \int |\varphi^*(y)\varphi^*(x-y)| dy &= \int |\varphi^*(xu)\varphi^*(x(1-u))| x du \\ &\leq \int_{|u|>1/3 \text{ and } |1-u|>1/3} k_3 |xu|^{-r} k_3 |x(1-u)|^{-r} |x| du \\ &\quad + \int_{|u|\leq 1/3} k_3^2 |x(1-u)|^{-r} |x| du + \int_{|1-u|\leq 1/3} k_3^2 |xu|^{-r} |x| du \\ &\leq k_3^2 3^r |x|^{1-2r} \int_{|u|>1/3} \frac{du}{|u|^r} + k_3^2 |x|^{1-r} \frac{2}{3} \left| \frac{3}{2} \right|^r + k_3^2 |x|^{1-r} \frac{2}{3} \left| \frac{3}{2} \right|^r \\ &\leq k_3^2 \left[\frac{2 \cdot 3^{2r-1}}{r-1} |x|^{1-2r} + 2^{2-r} 3^{r-1} |x|^{1-r} \right] \end{aligned}$$

Thus, if $|x| > 1$, $\int |\varphi^*(y)\varphi^*(x-y)| dy \leq C_r |x|^{1-r}$ and if $|x| \leq 1$, $\int |\varphi^*(y)\varphi^*(x-y)| dy \leq C_r$ with $C_r = k_3^2(2 \cdot 3^{2r-1}/(r-1) + 2^{2-r} 3^{r-1})$.

Lemma 5. *Let T_1, \dots, T_n be independent random variables and*

$$\nu_n(r) = (1/n) \sum_{i=1}^n [r(T_i) - \mathbb{E}(r(T_i))],$$

for r belonging to a countable class \mathcal{R} of measurable functions. Then, for $\epsilon > 0$,

$$(19) \quad \mathbb{E}[\sup_{r \in \mathcal{R}} |\nu_n(r)|^2 - 6H^2]_+ \leq C \left(\frac{v}{n} e^{-k_1 \frac{nH^2}{v}} + \frac{M_1^2}{n^2} e^{-k_2 \frac{nH}{M_1}} \right)$$

with $k_1 = 1/6$, $k_2 = 1/(21\sqrt{2})$ and C a universal constant and where

$$\sup_{r \in \mathcal{R}} \|r\|_\infty \leq M_1, \quad \mathbb{E} \left(\sup_{r \in \mathcal{R}} |\nu_n(r)| \right) \leq H, \quad \sup_{r \in \mathcal{R}} \frac{1}{n} \sum_{i=1}^n \text{Var}(r(T_i)) \leq v.$$

Usual density arguments allow to use this result with non-countable class of functions \mathcal{R} .

Proof of Lemma 5: We apply the Talagrand concentration inequality given in Klein and Rio (2005) to the functions $s^i(x) = r(x) - \mathbb{E}(r(T_i))$ and we obtain

$$P(\sup_{r \in \mathcal{R}} |\nu_n(r)| \geq H + \lambda) \leq \exp \left(-\frac{n\lambda^2}{2(v + 4HM_1) + 6M_1\lambda} \right).$$

Then we modify this inequality following Birgé and Massart (1998) Corollary 2 p.354. It gives

$$P(\sup_{r \in \mathcal{R}} |\nu_n(r)| \geq (1 + \eta)H + \lambda) \leq \exp \left(-\frac{n}{3} \min \left(\frac{\lambda^2}{2v}, \frac{\min(\eta, 1)\lambda}{7M_1} \right) \right).$$

To find inequality (19) we use the formula $\mathbb{E}[X]_+ = \int_0^\infty P(X \geq t)dt$ with $X = \sup_{r \in \mathcal{R}} |\nu_n(r)|^2 - 6H^2$.

Lemma 6. (*Viennet (1997)*) *Let (T_i) a strictly stationary process with β -mixing coefficients β_k . Then there exists a function b such that*

$$\mathbb{E}[b(T_1)] \leq \sum_k \beta_k \quad \text{and} \quad \mathbb{E}[b^2(T_1)] \leq 2 \sum_k (k+1)\beta_k$$

and for all function ψ (such that $\mathbb{E}[\psi^2(T_1)] < \infty$) and for all N

$$\text{Var} \left(\sum_{i=1}^N \psi(T_i) \right) \leq 4N \mathbb{E}[|\psi|^2(T_1)b(T_1)].$$

In particular, for functions (ψ_λ) , $\sum_\lambda \text{Var}(\sum_{i=1}^N \psi_\lambda(T_i)) \leq 4N(\sum_k \beta_k) \|\sum_\lambda |\psi_\lambda|^2\|_\infty$.

ACKNOWLEDGEMENTS

I would like to thank Fabienne Comte for her carefully readings of this work and Erwan Le Pennec for his help concerning the simulations.

REFERENCES

- Bakry, D., Milhaud, X., and Vandekerckhove, P. (1997). Statistique de chaînes de Markov cachées à espace d'états fini. Le cas non stationnaire. *C. R. Acad. Sci. Paris Sér. I Math.*, 325(2):203–206.
- Baraud, Y., Comte, F., and Viennet, G. (2001). Adaptive estimation in autoregression or β -mixing regression via model selection. *Ann. Statist.*, 29(3):839–875.
- Barron, A., Birgé, L., and Massart, P. (1999). Risk bounds for model selection via penalization. *Probab. Theory Related Fields*, 113(3):301–413.
- Baum, L. E. and Petrie, T. (1966). Statistical inference for probabilistic functions of finite state Markov chains. *Ann. Math. Statist.*, 37:1554–1563.
- Bickel, P. J., Ritov, Y., and Rydén, T. (1998). Asymptotic normality of the maximum-likelihood estimator for general hidden Markov models. *Ann. Statist.*, 26(4):1614–1635.
- Birgé, L. and Massart, P. (1998). Minimum contrast estimators on sieves: exponential bounds and rates of convergence. *Bernoulli*, 4(3):329–375.
- Butucea, C. (2004). Deconvolution of supersmooth densities with smooth noise. *Canad. J. Statist.*, 32(2):181–192.
- Cappé, O., Moulines, E., and Rydén, T. (2005). *Inference in hidden Markov models*. Springer Series in Statistics. Springer, New York.
- Chaleyat-Maurel, M. and Genon-Catalot, V. (2006). Computable infinite dimensional filters with applications to discretized diffusions. *Stochastic Process. Appl.*, 116(10):1447–1467.
- Cléménçon, S. (2000). Adaptive estimation of the transition density of a regular Markov chain. *Math. Methods Statist.*, 9(4):323–357.

- Cléménçon, S. (2003). Nonparametric estimation for some specific classes of hidden markov models. *Preprint Modal'X n° 03-9*: http://www.u-paris10.fr/65897276/0/fiche___pagelibre/.
- Cohen, A., Daubechies, I., and Vial, P. (1993). Wavelets on the interval and fast wavelet transforms. *Appl. Comput. Harmon. Anal.*, 1(1):54–81.
- Comte, F., Rozenholc, Y., and Taupin, M.-L. (2006). Penalized contrast estimator for adaptive density deconvolution. *Canad. J. Statist.*, 34(3):431–452.
- Douc, R., Moulines, É., and Rydén, T. (2004). Asymptotic properties of the maximum likelihood estimator in autoregressive models with Markov regime. *Ann. Statist.*, 32(5):2254–2304.
- Doukhan, P. (1994). *Mixing. Properties and examples*, volume 85 of *Lecture Notes in Statistics*. Springer-Verlag, New York.
- Fan, J. (1993). Adaptively local one-dimensional subproblems with application to a deconvolution problem. *Ann. Statist.*, 21(2):600–610.
- Jensen, J. L. and Petersen, N. V. (1999). Asymptotic normality of the maximum likelihood estimator in state space models. *Ann. Statist.*, 27(2):514–535.
- Klein, T. and Rio, E. (2005). Concentration around the mean for maxima of empirical processes. *Ann. Probab.*, 33(3):1060–1077.
- Koo, J.-Y. and Lee, K.-W. (1998). B -spline estimation of regression functions with errors in variable. *Statist. Probab. Lett.*, 40(1):57–66.
- Lacour, C. (2007a). Adaptive estimation of the transition density of a Markov chain. *Ann. Inst. H. Poincaré Probab. Statist.* To appear. Available online.
- Lacour, C. (2007b). Adaptive estimation of the transition density of a particular hidden Markov chain. *J. Multivariate Anal.* To appear. Available online.
- Leroux, B. G. (1992). Maximum-likelihood estimation for hidden Markov models. *Stochastic Process. Appl.*, 40(1):127–143.
- Masry, E. (1993). Strong consistency and rates for deconvolution of multivariate densities of stationary processes. *Stochastic Process. Appl.*, 47(1):53–74.
- Meyer, Y. (1990). *Ondelettes et opérateurs. I*. Hermann, Paris.
- Pensky, M. and Vidakovic, B. (1999). Adaptive wavelet estimator for nonparametric density deconvolution. *Ann. Statist.*, 27(6):2033–2053.
- Stefanski, L. A. (1990). Rates of convergence of some estimators in a class of deconvolution problems. *Statist. Probab. Lett.*, 9(3):229–235.
- Viennet, G. (1997). Inequalities for absolutely regular sequences: application to density estimation. *Probab. Theory Related Fields*, 107(4):467–492.
- Youndjé, É. and Wells, M. T. (2002). Least squares cross-validation for the kernel deconvolution density estimator. *C. R. Math. Acad. Sci. Paris*, 334(6):509–513.